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Tini a Tangaroa

A new model for the assessment of New Zealand rock lobster (Jasus edwardsii) stocks and an exploratory CRA 4 multi-area assessment New Zealand Fisheries Assessment Report 2018/53
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## EXECUTIVE SUMMARY

Webber, D.N.; Haist, V.; Starr, P.J.; Edwards, C.T.T. (2018). A new model for the assessment of New Zealand rock lobster (Jasus edwardsii) stocks and an exploratory multi-area CRA 4 assessment.

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This document describes the development of a new semi-generalised length structured stock assessment model for New Zealand red rock lobster (Jasus edwardsii) stock assessments, comparisons of this new model with the existing rock lobster assessment model, and an exploratory multi-area stock assessment for CRA 4.

The new model, named lobster stock dynamics (LSD), does not support puerulus randomisation trials or management procedure (MP) simulation yet so these aspects have not been included in this document. These features will be added to LSD this year. The LSD model was developed by D'Arcy Webber with input from the rock lobster stock assessment team contracted by the New Zealand Rock Lobster Industry Council Ltd. The comparisons between the single stock CRA 4 assessments done in the multi-stock length-based model (MSLM) and LSD were done by D'Arcy Webber and Charles Edwards. The exploratory multi-stock assessment for CRA 4 was done by Vivian Haist in the MSLM and by D'Arcy Webber and Charles Edwards in LSD. Paul Starr provided data for both comparisons.

The Rock Lobster Fishery Assessment Working Group oversaw this work: data files and all technical decisions were agreed beforehand or subsequently approved (and sometimes changed) by that group. Models were fit to CPUE indices, size frequency data, puerulus index data, and tag-recapture data. This document does not describe the procedures used to find acceptable model fits, instead models with acceptable model fits that were developed in the official CRA 4 stock assessment (Breen et al. 2017) were used as a starting point for the models presented in this document.

Due to its speed increases and additional flexibility, LSD will be an excellent platform for producing single-area and multi-area rock lobster stock assessments in the future. The multi-area modelling approach that was explored shows great promise. Not only does the approach have the potential to provide a greater understanding of what is happening at finer spatial scales, it is also more than capable of providing the status quo - that is stock status as a whole.

## 1. INTRODUCTION

This work addresses Objective 4 of the Ministry for Primary Industries (MPI) contract CRA2015-01A. This three-year contract, which began in April 2016, was awarded to the New Zealand Rock Lobster Industry Council Ltd. (NZ RLIC Ltd.), who sub-contracted Objective 4 to the authors of this report.

## Objective 4 - Stock assessment: To estimate biomass and sustainable yields for rock lobster stocks

This document presents auxiliary work contributing to this objective, including a new stock assessment model and an exploratory multi-area stock assessment analysis. The development of the new stock assessment model was to be done over two years. This document describes the work carried out during the first year of development and 2017 will see the completion of the new model (including puerulus randomisation code, management procedure simulation software, plotting routines and a user interface). The exploratory multi-area stock assessment is ongoing developmental work that may provide an alternative to standard single-area stock assessments in the future.

The National Rock Lobster Management Group (NRLMG) decided that the CRA 4 stock should be assessed in 2016. Data were compiled by a team comprising Paul Starr, D’Arcy Webber, and Paul Breen. See Starr et al. (2017) for the data preparation for the single area model. CRA 4 was assessed in the usual way, assuming a single homogeneous stock, using the purpose-built multi-stock length-based model (MSLM) of Haist et al. (2009); this work was done by Paul Breen, Paul Starr and Vivian Haist with input from D’Arcy Webber and Charles Edwards (see Breen et al. 2017).

During 2016, a new length structured model with similar dynamics to the MSLM was developed by D'Arcy Webber. The aim of this new model was to create software written in a modern programming language that can easily be upgraded (as needed) in the future, to decrease the time required for Bayesian inference and to add several new features or options to the assessment model. The new model was written using the Stan modelling language (Stan Development Team, 2016a). Stan is a probabilistic programming language for statistical inference written in $\mathrm{C}++$. The Stan language is used to specify a (Bayesian) statistical model with an imperative program calculating the log probability density function. Stan is licensed under the New BSD License and is named in honour of Stanislaw Ulam, pioneer of the Monte Carlo method. Stan implements gradient-based Markov chain Monte Carlo (MCMC) algorithms for Bayesian inference, stochastic, gradient-based variational Bayesian methods for approximate Bayesian inference, and gradient-based optimisation for penalized maximum likelihood estimation.

It is important that any new model/software be validated using some benchmark. Therefore, the new model was fit to the same data as MSLM, and results compared. The first comparison was done assuming a single homogeneous stock as described by Breen et al. (2017). At the same time, an experimental multi-stock assessment of CRA 4 was conducted using the multi-stock capability of MSLM and the new Stan model. Thus the models tested and presented in this document include:

- Single stock models using MSLM
- Single stock models using the new model
- Multi-stock models using MSLM
- Multi-stock models using the new model

Decisions on data and modelling choices were discussed and approved by the Rock Lobster Fishery Assessment Working Group (RLFAWG).

In the most recent stock assessment of CRA 4 in 2016, Breen et al. (2017) described the stock assessment and management procedure (MP) simulations, generated from the MSLM model. That model was fitted to tag-recapture data, standardised CPUE, historical catch rate data, length frequency data from voluntary logbooks and observer catch sampling, and puerulus settlement data. Changes in MLS and changes in selectivity caused by escape gap regulations were taken into account. Data for this work are described by Starr et al. (2017). The stock assessment was done in a workshop in

Wellington from 19 September through to 20 October; it was presented to the Mid-year Plenary on 1 November.

The present document describes the development of a new stock assessment model and some experimental multi-stock assessment models for CRA 4. This document does not provide any interpretation of the single-area CRA 4 stock assessment, which is done in Breen et al. (2017). A list of acronyms used throughout this document is provided in Table 1.

### 1.1 CRA 4

The CRA 4 (Figure 1) fishery extends from the Wairoa River on the east coast southwards along the Hawkes Bay, Wairarapa and Wellington coasts, through Cook Strait and north to the Manawatu River in the South Taranaki Bight. The CRA 4 total allowable catch (TAC) for 2016-17 was 592 t. Allowances set by the Minister of Fisheries were 35 t for customary catch, 85 t for recreational catch, 75 t for illegal unreported removals and a 397 t total allowable commercial catch (TACC). The CRA 4 commercial fishery is open all year. The minimum legal size (MLS) is 54 mm tail width (TW) for males and 60 mm TW for females for both the commercial and recreational fisheries.


Figure 1: The CRA 4 Quota Management Area (QMA) and its statistical areas (912, 913, 914, 915 and 934).

The CRA 4 commercial fleet comprised 50 vessels in the 2016-17 fishing year ${ }^{1}$ (Starr 2016). Most vessels in the fleet operate from coastal bases in isolated rural areas on the Hawkes Bay and Wairarapa

[^0]coastlines. The CRA 4 commercial catch supports several processing and export operations in Napier, Wellington and Auckland.

Potting and hand gathering are the preferred methods for recreational fishers in this area. As in most CRA areas, the majority of the recreational catch is taken in the summer months. The region also sustains a recreational fishing and dive charter industry during summer. Lobsters are very important to Maori in this area, and the customary allowance allows lobsters to be taken under permit for use by the marae. This is a trap or pot fishery, conducted by small boats on day trips, fishing in relatively shallow waters.

The stock assessment and data preparation separate the autumn-winter (AW, April through to September) and spring-summer (SS) seasons. The stock is managed with an operational management procedure (MP) that determines the TACC, the primary management tool. Allowances are added by the Minister for the non-commercial fisheries to produce a TAC. Other management measures include protection of ovigerous (berried) females, MLS by sex, and escape gaps in pots.

## 2. THE ROCK LOBSTER STOCK ASSESSMENT MODELS

Rock lobster assessment models (and invertebrate models in general) are typically length based because invertebrates are difficult and expensive to age, rendering the collection of age-based data infeasible. The central component of length based models is the growth model and variation within that model, which describes the transition of individuals between length classes with each time step. Modern assessment models are integrated (see Maunder \& Punt 2013; Punt et al. 2013), meaning that they make use of a variety of data sources to estimate component parameters. For length based models, mark-recapture-at-length data provide the primary information for estimation of the length transition process. Length frequency data, when tracked over time, are also useful for estimation of growth, and provide additional information on the fishery selectivity. Finally, catch and abundance information (in this case the CPUE), inform the estimation of stock productivity via natural mortality and stock recruitment. Based on these fundamental ideas concerning a length based integrated approach, we describe the current New Zealand stock assessment approach for rock lobster, along with improvements being made during development of a new model.

### 2.1 Multi-stock length-based model (MSLM)

The Bayesian multi-stock length-based model (MSLM) was described by Haist et al. (2009). The model is implemented in AD Model Builder (ADMB, Fournier et al. 2012). The model time step is specified and the length of the time step can vary during the period being simulated. The model's number and width of size bins is specified. Fishing is modelled by taking into account the observed catch, the MLS that can change during the period simulated, estimated seasonal vulnerability, and estimated size-selectivity of the fishing gear that can vary over time. The model fits the catch that is limited by MLS and a restriction on landing ovigerous females (SL catch), comprising the commercial and recreational catches, and separately fits the catch not limited by these regulations (NSL catch), comprising the illegal and customary catches, which are assumed to take all the lobsters caught by a pot.

Differences in the growth rate between males and females justifies the two sexes being represented separately in the model. In addition, because of the restriction against landing berried females, the female partition is further divided into mature and immature individuals. The model therefore tracks three components of the population, and at each time step, the number of male, immature female and mature female lobsters in each size class is updated as a result of somatic growth and annual recruitment to the model. Recruitment occurs to a specified mean size with specified size variation and can vary over time. Somatic growth can be divided into distinct "epochs". Natural mortality is estimated but assumed to be constant over time, sizes and sexes. Handling mortality of returned lobsters (undersized and berried females) is assumed and constant.

A growth transition matrix, based on estimated sex-specific growth parameters, specifies the probability of an individual lobster remaining in the same size bin or growing into each of the other size bins, including smaller ones. Maturation of females is described by a two-parameter logistic curve.

The model calculates biomass vulnerable to the fishery at each time step from numbers-at-size for each sex, the size-weight relationship, the female maturity (for the SL fishery, mature females are assumed to be berried and thus not legal in the AW season), the MLS (for the SL fishery only), the sex-specific trap selectivity-at-size and the sex-specific seasonal vulnerability. MLS has changed over time and is input as a covariate for each year.

The model is fit to abundance indices, size composition data, tag-recapture data and puerulus settlement data. The model can be fit to these data using penalised maximum likelihood or Markov chain Monte Carlo simulations (MCMC). Although Bayesian procedures are time-consuming, they are recommended to be the default method for estimating uncertainty in stock assessments (Magnusson et al. 2012).

### 2.2 Lobster stock dynamics (LSD) model

The new Bayesian multi-stock length-based model has been named lobster stock dynamics (LSD). Like the MSLM, LSD is an integrated model (see Maunder \& Punt 2013; Punt et al. 2013) that estimates most structural parameters by fitting to several data sets simultaneously. However, LSD was written using the state of the art Stan modelling language making use of its very efficient Hamiltonian Monte Carlo (HMC) sampler (Stan Development Team, 2016b). The advantages of HMC and Stan are outlined in Monnahan et al. (2016). Although Stan is a relatively young language compared to the ADMB suite of programs, and therefore has fewer features and functions, there are many good reasons to use Stan. Stan has a much broader user group than ADMB, was built by Bayesian statisticians to do MCMC (yet it can also do optimisation and variational Bayes), deals with all parameter transformations and Jacobians automatically, and makes running multiple MCMC chains on different computer cores very easy.

There were several reasons for wanting to recode the MSLM described by Haist et al. (2009). Most importantly, recoding the assessment provided an opportunity to review and potentially improve the existing rock lobster stock assessment model. New ideas or features were often incorporated into the code during each year that the MSLM was used. While these new features were extensively tested before being used for stock assessment, the additional code was often added without much consideration for efficiency due to time constraints. Therefore, as the software ages and additional features are "bolted on", the code can become cumbersome and slow.

The new model was redeveloped from the ground up to be efficient and (strictly) Bayesian. The plan is to develop LSD over 2 years. The first year saw the model developed and used alongside the MSLM during the CRA 4 workshop in 2016. The goal was to keep as much as possible the same between the two models during the first year so that we could compare them side-by-side. The second year (2017) will see the puerulus randomisation and MP simulation code developed and further changes to the model code. Therefore, the models will diverge somewhat during code development in 2017.

Improvements to code structure that have already been implemented include:

- Parameter mapping - the desired number of most parameters is first specified, then these parameters are mapped to the model by specifying, for example, the area, sex, and year to which the parameter is relevant. For example, the desired number of each of the selectivity parameters is specified, then each selectivity parameter is mapped to an area, sex, and year. This means the user could share selectivity parameters across areas or sexes, or change selectivity by year. While much more flexible, the new approach can mimic the current assessment easily. This same method has been rolled out to most parameters in the model (e.g., vulnerability, natural mortality, growth).
- Speed - Bayesian inference using MCMC can be slow, especially for multi-area stock assessment models. Therefore, LSD was written with computer code efficiency (speed) in mind. The LSD code is split into different modules including the base model code, parameters, priors, and various functions. When the user begins a model run only those modules that are required are gathered up and compiled into the final model. This excludes any unutilised code and helps reduce the computational workload. Code written in this way is also easier to error check and add to at a later date because the code is split up into logical chunks that can be worked on independently.
- Automation - reducing the workload for the user was another key goal when writing LSD. A user interface is being developed that will allow the user to import and groom assessment data, change model settings, and run models. Automating some of the steps involved in stock assessment can reduce the risk of user error during the stock assessment workshop. Until the user interface is finished, makefiles have been developed and these can be used to carry out the same tasks easily from the command line.

Finally, the process for conducting model runs and doing a stock assessment differs between MSLM and LSD. Like MSLM, LSD is controlled using a set of input files that contain the model specifications, data, and initial parameter values. However, LSD uses makefiles to do model runs via the CmdStan interface to Stan (Stan Development Team, 2016a). The makefiles can also be used to produce model outputs (including plots and tables) and even do MCMC remotely on the cloud.

## 3. SINGLE-AREA MODEL COMPARISONS

This section compares single-area model runs developed using the two different stock assessment models: the MSLM model written in ADMB; and the LSD model written in Stan. The objective was to ensure that the two models were satisfactorily similar (a perfect match would not be reasonable given they are written in different software and run on different computers/operating systems with different machine precisions).

The structure of the model run used for the comparison loosely followed the base case of the Breen et al. (2017) CRA 4 model. A great deal of effort went into trying to match the two models as closely as possible. To ensure that the two models produced as similar results as possible three different comparisons were made:

1. An MPD fit was obtained using the MSLM. Estimated parameter outputs, plus all of the model inputs (data, likelihood weights, fixed parameters and estimated parameters) were then used to do a single fixed parameter model run in LSD. The model fits to the data, derived quantities, SDNRs and MARs, and likelihood components were then compared.
2. The initial parameter values used to fit the MSLM were used to initialise an MPD fit in Stan. All of the other model inputs (data, likelihood weights, and fixed parameters) were the same in the two models. The model fits to the data, derived quantities, SDNRs and MARs, and likelihood components were then compared.
3. A Bayesian posterior was obtained using MCMC in both the MSLM model and the LSD model. In both models the initial values were the same. The posteriors were then compared.

More detail for each of these comparisons follows.

### 3.1 Fixed parameter comparisons

The MSLM model was fitted to the CRA 4 data providing the MPD, or the model's best estimate of each of the model's non-fixed parameters. Using these same parameter estimates as an input for LSD to do a single model run, along with the other fixed model inputs (data, likelihood weights, and fixed parameters) yielded very similar results (Table 2). For example, the prior contribution to the objective function in MSLM was -41.77 compared with -41.77009 in LSD. A difference this small is likely to be due to rounding by ADMB. The total objective function value was 8846.82 in MSLM compared with
8846.92 in LSD, again very close. Each of the individual components of the total likelihood (i.e. CPUE, CR, puerulus index, tags, and length-frequency) were also very similar. This was also true of several other model runs with different structures, including multi-area models (not presented here).

The SDNRs and MARs were also very similar (Table 2). For example, the CPUE SDNR was 1.208 and 1.219 in MSLM and LSD, respectively. The CPUE MAR was 0.827 and 0.854 in MSLM and LSD, respectively. The biggest difference was in the length-frequency SDNR and MAR values where the SDNR was 1.047 and 0.867 , and the MAR was 0.181 and 0.055 , in MSLM and LSD, respectively. These differences, although relatively small, may arise from different methods for calculating the median in ADMB and Stan and require further investigation.

Several comparisons were also made between derived quantities (e.g., MSY, $\mathrm{B}_{\text {MSY }}, \mathrm{B}_{\text {ref }}$ ) calculated in each model. These were all very similar as well (Table 2). The RL FAWG agreed that the fixed parameter comparisons were adequate and that LSD provides a very close match to the MSLM.

### 3.2 Penalised maximum likelihood comparisons

Optimisation involves finding the maximum likelihood (aka MPD, MAP, penalised maximum likelihood, etc), the set of parameter values that maximises the negative log-likelihood and provides the best fit to the data given the structure of the model. Both ADMB and Stan use gradient-based minimisers to find the optima. Before beginning any optimisation routine it is necessary to provide initial values for each of the model's estimated parameters, giving the optimisation algorithm a place to start from. It is desirable to choose sensible initial values that are reasonably close to the optimum parameter values because this usually avoids any numerical issues during optimisation and speeds up the algorithm.

In this experiment the same initial values that were provided to the MSLM model were used in the LSD model. We noted that Stan did take about $10 \%$ longer than ADMB to minimise, but found a slightly better optimum (-1337.6 for Stan compared with -1337.3 for ADMB). However, the convergence criterion might not have been directly comparable between the two models. Like ADMB, Stan has several optimisation options (i.e., algorithm, iter, obj_tol, tol_rel_obj, tol_grad, tol_rel_grad, tol_param, history_size) and tweaking these options can result in faster convergence at the expense of the accuracy of the optimum, or vice versa.

This process was repeated several times for various different models with the same outcome - ADMB and Stan finding approximately the same optimum with the same parameter estimates (implying the same derived quantities as demonstrated in section 3.1). The RL FAWG agreed that the penalised maximum likelihood comparisons were adequate and that Stan provides a very close match to the ADMB optima.

### 3.3 Markov chain Monte Carlo (MCMC) comparisons

The final check was to ensure that the different MCMC algorithms used by ADMB and LSD were comparable and resulted in similar posterior distributions. Again, this check was done using a singlearea model. An MCMC was run in MSLM with 1 million iterations, retaining every $1000^{\text {th }}$ sample, and took over 48 hours to complete. Similarly, an MCMC was run in LSD but different MCMC specifications were used because Stan's MCMC sampling algorithm is much more efficient than the Metropolis-Hastings MCMC used by ADMB. A total of 1000 samples from the posterior were obtained by specifying 4 chains, each of 500 iterations, thinning every second iteration. A burn-in (or warm-up) of 500 iterations was also done, but these samples were discarded. This took approximately 24 hours for Stan to complete.

The initial MSLM MCMC mixing was tolerable during the first couple of hundred samples before abruptly switching to a different (well mixed) parameter space (Figure 2). The same unusual behaviour was observed in the initial LSD MCMC. The posterior for this particular model appeared to
be bimodal in both MSLM and LSD. In LSD, chains 1 and 4 explored parameter space about one of the modes, chain 3 explored parameter space about the other mode, and chain 2 explored parameter space about the first mode, but abruptly switched modes about $70 \%$ of the way through the MCMC (Figure 3). These two posterior modes are clearly visible in Figure 4, especially for some of the growth parameters and $\mathrm{R}_{0}$. These model runs were consequently called the bimodal model runs.

The mode switching behaviour seems to be caused by switches in productivity (i.e., growth) and mortality in the stock (i.e., male and female growth rates increase, while natural mortality and $\mathrm{R}_{0}$ decrease) - both scenarios result in very similar objective function values (see $f$ in Figure 2A and lp_ in Figure 3A and Figure 4A). Multi-modality is not surprising in stock assessment models but is rarely documented. To stop this behaviour, and make comparisons easier, weak priors were placed on the growth parameter $\mathrm{G}_{\text {alpha, }}$, the $\mathrm{R}_{0}$ parameter, and each of the selectivity parameters. The Stan developers recommend the use of wide priors for all model parameters, unless we do have some prior knowledge of what the parameter value should be. We used wide normal distributions that are relatively flat across the credible range for each parameter (Figure 5). We then repeated the MCMC's outlined above using these new priors and called these the wideprior model runs.

In the wideprior model runs, both the MSLM MCMC and the LSD MCMC were well mixed (Figure 6 ) and no mode switching behaviours were observed (Figure 8). Figure 9 was created to compare the posterior distributions for several parameters and derived quantities for the MSLM model without the priors applied and the LSD with and without the priors applied (the ADMB model run with the priors was not completed when these figures were created). These figures suggest that the priors had little influence on the posterior distributions of most model parameters, except those parameters that were multi-modal. Overall, the posteriors produced using ADMB and Stan were very similar and result in almost identical stock assessment outcomes.

We note that each MCMC iteration is much slower in Stan, but not as many iterations are needed because mixing is much more efficient, which reduces the amount of thinning needed (e.g., in the LSD MCMCs outlined above we discarded every second iteration, rather than every hundredth or thousandth iteration in ADMB), resulting in a much faster MCMC runtime overall. It is also convenient that there is no need to start the MCMC near the MAP in Stan (but it does help to reduce the warm-up period), and positive definite Hessian (pdH) problems no longer apply because Stan approximates the covariance matrix during the warm-up phase. In fact we don't even need to optimise to run an MCMC in Stan. Because of its speed increases and additional flexibility, LSD will be an excellent platform for finer scale spatial modelling in the future.

## 4. EXPLORATORY MULTI-AREA MODEL

The CRA 4 multi-area model documented here is exploratory work that aims to eventually provide a framework for producing stock assessment outputs at a finer spatial scale than standard single-area stock assessments (e.g., estimates of stock size by statistical area or some combination of aggregated statistical areas). This work remains exploratory, for now, because of the number of choices and challenges that are faced when moving to multi-area models. For example, how many sub-areas should the model consider? Which statistical areas should be separate or combined? How many growth morphs (areas with different growth rates or time periods where growth differs) should be modelled? How many different selectivity curves should the model include (e.g., a different selectivity curve for each area, shared across some or all areas, one for every year)? How many natural mortality parameters should be used (one for each area, shared across areas, or some combination)? Determining best practices for these types of modelling decisions could take years of research. Despite this, we believe that, in the near future, multi-area models like this will augment the single-stock/single-area models that dominate fisheries stock assessment.

As with the single area models, fixed parameter tests and comparisons were done and the ADMB and Stan optimisers were tested. The fixed parameter tests resulted in very similar objective function values and prior contributions (Table 3). The optimisers made their way to the same place in ADMB and Stan. The RL FAWG agreed that these comparisons were adequate and that the model runs in

LSD and MSLM were sufficiently similar. We called these the spatial model runs. Because of the exploratory status of the spatial model runs we provide limited interpretation of model outputs in this document.

### 4.1 Multi-area model structure

CRA 4 is made up of five statistical areas (Figure 1) and it was desirable to model the stock at the finest spatial scale possible. The data used as inputs to the stock assessment are provided by statistical area, so this is currently the smallest spatial scale possible. However, summaries of the data suggested that limited data were available for statistical area 934. Therefore, 915 and 934 (adjacent statistical areas) were combined in the multi-area assessment model presented here, resulting in a four area model. These four areas will be referred to as 912, 913, 914 and $915+934$. Combining two statistical areas in this way assumes that the dynamics in these two areas are the same (e.g., selectivity, vulnerability, catchability, and catch rates), requires that the data be aggregated for these two areas (e.g., annual catches need to be summed), and provides stock assessment outputs for both areas combined (e.g., the reference biomass provided by the model for $915+934$ cannot be split into its constituent statistical areas).

Growth was not assumed to be consistent across all four modelled areas; instead, three different growth areas were defined. Growth in area 912 was estimated from individuals tagged in 911 and 912 (auxiliary data from CRA 3 were introduced because there were few recaptures in 912). Growth in areas 913 and 914 was assumed to be the same and was estimated using tag-recaptures from these two areas combined. Growth in $915+934$ was estimated using tag-recaptures from 915 (there were no recaptures reported for 934). These area splits were chosen based on careful consideration of the amount of tag-recapture data available for each statistical area, geographical proximity of the areas, and inspection of the tag data residuals after fitting an initial exploratory model (largely through trial and error).

Over 30 different spatial model runs were done in MSLM and LSD. Several different model structures were explored during these runs including:

- 2-sex models
- 3-sex models
- Dropping logbook (LB) data
- Various different selectivity assumptions
- Various different vulnerability assumptions
- Various different natural mortality assumptions


### 4.2 Data

Data for the single area implementation of the CRA 4 stock assessment model are described by Starr et al. (2017). However, the multi-area implementation of the CRA 4 stock assessment required most of the data to be area-specific. The preparation of these data is described in Appendix A.

### 4.3 Results

We present the posterior distributions for just one of the model runs discussed above because time was limited during the assessment workshop and MCMC sampling was slow for these multi-area models. Using the LSD model, a total of 1000 samples from the posterior were obtained by specifying 4 MCMC chains, each of 500 iterations, thinning every second iteration. A burn-in or warm-up of 250 iterations was also done, but these samples were discarded. This took approximately 3 days for Stan to complete. The MCMC for the multi-stock model coded in MSLM was too slow to get adequate posteriors within a practical timeframe (Figure 10). The LSD MCMC's were well mixed within this timeframe (traces are shown in Figure 11 and histograms in Figure 12).

The area-specific model fits to the CPUE and CR series were better in some areas than others (Figure 13). The fit to the CR data is excellent in all statistical areas (Figure 13). In 912, the fit to the CPUE data is good until about 2010 where the model begins to stray from the data in both the AW and SS. In 913 and 914, the fit to the CPUE is excellent. In the last model area that consists of statistical areas 915 and 934 combined, the fit to the CPUE data is reasonable given the much more volatile CPUE series.

Fits to length-frequency (LF) data were generally very good and often better than in the single-area models (examples of these fits are provided in Figure 14). This improvement in fits to the LF data is due to different selectivity curves being estimated among areas in the model. Specifically, selectivity was parameterised to be logistic and sex-specific in all areas, and was assumed to be the same in areas 913 and 914, with different selectivities for areas 912 and 915+934 (Figure 15).

Recruitment deviations were estimated independently in each area. Despite this, many of the peaks and troughs in recruitment occurred during the same years in each area (Figure 16). Recruitment was generally higher in areas 913 and 914 . In all areas recruitment was estimated to be relatively low in recent years.

Reference biomass in areas 912 and 913 was similar throughout the stock assessment timeframe, starting from approximately 0.375 tonnes and declining until the mid-1950s, followed by an increase up to the mid-1960s, then declining again until the mid-1990s (Figure 17). From the mid-1990s to about the year 2000 the reference biomass increased again in all areas, but has since declined again except in area 915+934 where the reference biomass increased up until 2014.

## 5. DISCUSSION

Over the past decade, the MSLM (described by Haist et al. 2009) has served as the foundation of all stock assessments for New Zealand rock lobster. The code for this model was written in ADMB, a platform that has become increasingly outdated and unsuitable for further development of the New Zealand rock lobster stock assessments. New platforms that outperform ADMB computationally have recently become available and the rock lobster team has selected Stan for development of future model code.

Stan, and therefore the new LSD model, has several benefits over ADMB. One of the most important benefits is that the Hessian matrix does not need to be positive definite in Stan for MCMC sampling to begin (this is a requirement in ADMB). This opens up a suite of models that can be taken to MCMC that were previously rejected because ADMB did not calculate a positive definite Hessian. The LSD model is also a good way forward for spatial modelling because MCMC mixing is much faster. This is because Stan's Hamiltonian Monte Carlo (HMC) is a much more efficient MCMC sampler. Model structures that were once beyond reach because MCMC mixing was too slow to produce reasonable posteriors within a practical time frame are now possible using LSD.

The multi-area modelling approach that was explored during the CRA 4 stock assessment is most likely credible as a future option for most rock lobster stock assessments. Not only does the approach provide greater understanding of what is happening at the statistical area scale, it also provides the status quo - that is, stock status as a whole (produced by aggregating across modelled areas). Although the current management framework in New Zealand does not manage rock lobster at the statistical area scale, information at finer spatial scales is still useful. For example, fishers could voluntarily manage their catch within statistical areas, so long as they are compliant at the QMA level.

This work is still experimental because multi-area models are a lot more complex than single-stock models. There are many more modelling choices to make, some of these choices are structural and can have potentially large effects on stock assessment outcomes. Inference is much slower than comparable single-area models. Spatially explicit data are also more difficult to generate, often requiring strong assumptions about area splits when data are lacking. And the results are more time
consuming to interpret. For example, as the number of areas increases, so does the number of figures and tables required to interpret the results from each of these areas and all of these areas combined. This means that choosing what to present also becomes important. For these reasons, it is important that these models remain experimental until we are confident that good stock assessment inference is manageable within practical timeframes and that the tools to communicate the results effectively are well established.

The LSD framework as a whole is also currently incomplete. The remaining tasks to complete the LSD framework include:

- Further work on the model documentation, including a technical report with all of the model equations and a user manual - important for communication and to improve ease of use
- Development of the management procedure (MP) simulation code
- Development of the puerulus randomisation code
- Develop new procedures to groom length-frequency (LF) data and a new approach to weighting these data whereby effective number of samples are sex-specific (i.e., males, immature females and mature females have their own effective sample size). This will do away with the need for sex-specific data weighting and instead an overall LF data weight will be used (or this data weighting could be for catch sampling and logbook data sources).
- Include code and options for self-weighting LF distributions (i.e., Dirichlet, logistic normal) which have the potential to do away with iterative data weighting methods.
- May want to modify LSD in the future to deal with spatial structure better (e.g. random effects or CAR priors for parameters; fit data for each area or summed across areas to help deal with data being by statistical area from 1979 onwards).


## 6. ACKNOWLEDGEMENTS

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Table 1: Acronyms used throughout this document.
Acronym Meaning
ADMB AD - Model builder
AW Autumn-winter season in model
CAR Conditionally autoregressive (prior)
CPUE Catch per unit effort
HMC Hamiltonian Markov chain
LF Length-frequency
LSD Lobster stock dynamics model
MAR Median of the absolute residual
MCMC Markov chain Monte Carlo
MLS Minimum legal size (mm)
MP Management procedure
MPD Maximum posterior density
MPI Ministry for Primary Industries
MSLM Multi-stock length-based model
pdH Positive definite Hessian
QMA Quota Management Area
RLFAWG Fisheries New Zealand Rock Lobster Fishery Assessment Working Group
SDNR Standard deviation of the normalised residual
SS Spring-summer season in model
TAC Total allowable catch
TACC Total allowable commercial catch
TW Tail-width (mm)

Table 2: Comparisons between the MSLM model and the LSD model with a single-area model run. Values compared include the total objective function value, prior contribution, components of the loglikelihood (LL) with associated SDNR and MAR data weights, and several derived quantities.

| Comparison | MSLM | LSD |
| :--- | ---: | ---: |
| Total objective function | 8846.82 | 8846.92 |
| Prior | -41.77 | -41.77009 |
| LF SDNR | 1.04686 | 0.867319 |
| LF MAR | 0.180631 | 0.055466 |
| LF LL | 6484.28 | 6484.29 |
| Sex-ratio SDNR | 0.976597 | 0.978484 |
| Sex-ratio MAR | 0.419328 | 0.42027 |
| Puerulus SDNR | 1.05679 | 1.07144 |
| Puerulus MAR | 0.743375 | 0.742744 |
| Puerulus LL | -25.9415 | -25.9415 |
| CR SDNR | 0.816606 | 0.851262 |
| CR MAR | 0.591625 | 0.56894 |
| CR LL | -24.8143 | -24.83061 |
| CPUE SDNR | 1.20873 | 1.21852 |
| CPUE MAR | 0.826705 | 0.853768 |
| CPUE LL | -128.045 | -127.9446 |
| Tag SDNR | 1.14928 | 1.14954 |
| Tag MAR | 0.545004 | 0.545645 |
| Tag LL | 2577.16 | 2577.14 |
| MSY | 673.877 | 673.575 |
| B $_{\text {MSY }}$ | 245.201 | 245.187 |
| B $_{\text {ref }}$ | 494.009 | 492.41 |
| B $_{2016} /$ B $_{\text {ref }}$ | 0.768529 | 0.7693345 |
| B $_{2016} /$ B $_{\text {MSY }}$ | 1.54836 | 1.545057 |
| F mult | 3.53 | 3.52 |

Table 3: Comparisons between the MSLM model and the LSD model with a multi-area model run. Values compared include the total objective function value, prior contribution, components of the log-likelihood (LL) with associated SDNR and MAR data weights, and several derived quantities.

| Comparison | MSLM | LSD |
| :--- | ---: | ---: |
| Total objective function | 13377.3 | 13380.8 |
| Prior | -82.861 | -82.861 |
| Tag prior | 12.487 | 12.487 |
| LF LL | 11667.9 | 11667.3 |
| CR LL | -73.34 | -71.48 |
| CPUE LL | -381.14 | -378.86 |
| Tag LL | 2234.27 | 2234.26 |



Figure 2A: MCMC trace plots of the likelihood profile (f) and model parameters from the bimodal model using the MSLM model in AMDB.


Figure 2B: MCMC trace plots of model parameters from the bimodal model using the MSLM model in AMDB.


Figure 2C: MCMC trace plots of model parameters and derived quantities from the bimodal model using the MSLM model in AMDB.


Figure 2D: MCMC trace plots of derived quantities from the bimodal model using the MSLM model in AMDB.


Figure 3A: MCMC trace plots of the likelihood profile (lp__) and model parameters from the bimodal model using the LSD model in Stan.


Figure 3B: MCMC trace plots of model parameters from the bimodal model using the LSD model in Stan.


Figure 3C: MCMC trace plots of model parameters and derived quantities from the bimodal model using the LSD model in Stan.


Figure 3D: MCMC trace plots of derived quantities from the bimodal model using the LSD model in Stan.


Figure 4A: MCMC histograms of the likelihood profile (lp__) and model parameters from the bimodal model using the LSD model in Stan.


Figure 4B: MCMC histograms of several of the model parameters from the bimodal model using the LSD model in Stan.


Figure 4C: MCMC histograms of several of the model parameters and derived quantities from the bimodal model using the LSD model in Stan.


Figure 4D: MCMC histograms of several derived quantities from the bimodal model using the LSD model in Stan.


Figure 5: The priors placed on the model parameters $\mathrm{R}_{0}$, $\mathrm{G}_{\text {alpha, }}$ sel50 and sel95 in the wideprior model runs using both the MSLM model in AMDB and the LSD model in Stan.


Figure 6A: MCMC trace plots of the likelihood profile (f) and model parameters from the wideprion model using the MSLM model in AMDB.


Figure 6B: MCMC trace plots of model parameters from the wideprior model using the MSLM model in AMDB.


Figure 6C: MCMC trace plots of model parameters and derived quantities from the wideprior model using the MSLM model in AMDB.


Figure 6D: MCMC trace plots of derived quantities from the wideprior model using the MSLM model in AMDB.


Figure 7A: MCMC trace plots of the likelihood profile (lp_) and model parameters from the wideprior model using the LSD model in Stan.


Figure 7B: MCMC trace plots of model parameters from the wideprior model using the LSD model in Stan.


Figure 7C: MCMC trace plots of model parameters and derived quantities from the wideprior model using the LSD model in Stan.


Figure 7D: MCMC trace plots of derived quantities from the wideprior model using the LSD model in Stan.


Figure 8A: MCMC histograms of the likelihood profile (lp__) and model parameters from the wideprior model using the LSD model in Stan.


Figure 8B: MCMC histograms of model parameters from the wideprior model using the LSD model in Stan.


Figure 8C: MCMC histograms of model parameters and derived quantities from the wideprior model using the LSD model in Stan.


Figure 8D: MCMC histograms of derived quantities from the wideprior model using the LSD model in Stan.


Figure 9A: MCMC densities of the likelihood profile ( $\mathrm{lp} \_$_) and model parameters comparing the MSLM model (ADMB uniform in figure), bimodal model (Stan uniform in figure), and the wideprior model (Stan normal in figure).




Model



| $\square$ |
| :--- |
| ADMB uniform |
| $\square$ |
| STAN normal |
| STAN uniform |



Value

Figure 9B: MCMC densities of model parameters comparing the MSLM model (ADMB uniform in figure), bimodal model (STAN uniform in figure), and the wideprior model (STAN normal in figure).


Figure 9C: MCMC densities of model parameters comparing the MSLM model (ADMB uniform in figure), bimodal model (STAN uniform in figure), and the wideprior model (STAN normal in figure).


Figure 9D: MCMC densities of model parameters and derived quantities comparing the MSLM model (ADMB uniform in figure), bimodal model (STAN uniform in figure), and the wideprior model (STAN normal in figure).


Figure 10A: MCMC trace plots of the likelihood profile (f) and model parameters from the MSLM spatial model.


Figure 10B: MCMC trace plots of model parameters from the MSLM spatial model.


Figure 10C: MCMC trace plots of model parameters and derived quantities from the MSLM spatial model.


Figure 10D: MCMC trace plots of model parameters and derived quantities from the MSLM spatial model.


Figure 10E: MCMC trace plots of derived quantities from the MSLM spatial model.


Figure 11A: MCMC trace plots of the likelihood profile ( $\mathrm{lp} \_$) and model parameters from the LSD spatial model.


Chain
$\begin{aligned} & -1 \\ & -2 \\ & -3 \\ & -3 \\ & -4\end{aligned}$

Figure 11B: MCMC trace plots of model parameters from the LSD spatial model.


Figure 11C: MCMC trace plots of model parameters and derived quantities from the LSD spatial model.



Figure 11E: MCMC trace plots of derived quantities from the LSD spatial model.


Figure 11F: MCMC trace plots of derived quantities from the LSD spatial model.


Figure 12A: MCMC histograms of the likelihood profile (lp__) and model parameters from the LSD spatial model.


Figure 12B: MCMC histograms of model parameters from the LSD spatial model.


Chain

| 1 |
| :---: |
| 2 |
| 3 |
| 3 |
|  |

Figure 12C: MCMC histograms of model parameters and derived quantities from the LSD spatial model.


Figure 12D: MCMC histograms of derived quantities from the LSD spatial model.


Figure 12E: MCMC histograms of derived quantities from the LSD spatial model.


Figure 12F: MCMC histograms of derived quantities from the LSD spatial model.


Figure 12G: MCMC histograms of derived quantities from the LSD spatial model.


Figure 13: Posteriors of the LSD model fit to catch per unit effort (CPUE) and catch rate (CR) by area and season (AW=autumn-winter, SS=spring-summer) from the LSD spatial MCMC. The shaded areas show the $\mathbf{5 \%}, \mathbf{2 5 \%}, \mathbf{5 0 \%}, \mathbf{7 5 \%}$ and $\mathbf{9 5 \%}$ quantiles of the posterior; error bars about the CPUE and CR data are one standard deviation.


Figure 14A: Posteriors of the LSD model fit to LF data from SS 1986 to AW 1991 by area, year, season (AW=autumn-winter, SS=spring-summer), and data source (CS=catch sampling, LB=logbook) in the LSD spatial MCMC. The shaded areas show the $5 \%, 25 \%, 50 \%, 75 \%$ and $95 \%$ quantiles of the posterior (because the posterior is so tight it is often difficult to discern the shaded quantiles); the vertical dashed black lines show the lower limit of the data that is fitted in the model, the MLS, and the upper limit; the value " $n$ " shown on each panel is the effective sample size; the value " $N$ " is the total number of individuals measured.


Figure 14B: Posteriors of the LSD model fit to LF data from AW 2013 to AW 2015 by area, year, season, and data source in the LSD spatial MCMC.


Figure 15: Posterior of the selectivity by sex and area in the LSD spatial MCMC. Shaded areas show the $5 \%, 25 \%, 75 \%$ and $95 \%$ quantiles of the posterior; the heavy solid line is the median of the posterior distribution; the dashed line is the MPD.


Figure 16: Posterior trajectory of recruitment (millions of individuals) to the model from 1945-2015 and projected recruits from 2016-2019 from the LSD spatial MCMC. The black shaded areas show the 5\%, $\mathbf{2 5 \%}, \mathbf{5 0 \%}, \mathbf{7 5 \%}$ and $95 \%$ quantiles of the recruitment posterior; the green shaded areas show the $5 \%$, $\mathbf{2 5 \%}, \mathbf{5 0 \%}, \mathbf{7 5 \%}$ and $\mathbf{9 5 \%}$ quantiles of the $\mathrm{R}_{0}$ posterior; the dashed red lines show the MPD and the MPD for $R_{0}$; the dashed vertical black line shows 2015, the final fishing year of the model reconstruction.


Figure 17: Vulnerable biomass trajectory from 1945-2019 by season and area [top four panels] and aggregated across all areas [bottom panel] in the LSD spatial MCMC. Shaded areas show the $90 \%$ credibility intervals; the heavy solid line is the median of the posterior distributions; the dashed line shows the MPD; the vertical line shows 2015, the final fishing year of the model reconstruction. Biomass before 1979 is annual and plotted using the YR coding.

## APPENDIX A. Data Preparation

## A. 1 Introduction

This document describes the catch and CPUE data assembled for use in the 2016 CRA 4 multi-area rock lobster (Jasus edwardsii) stock assessments, defined by individual statistical areas (see Table A. 1 below). It also describes the biological length frequency and tagging data assembled for the same multi-area model.

Table A.1: Sub-stock definitions for CRA 4 multi-area stock assessment, showing the rock lobster statistical area definitions used.

## Sub-stock name Statistical area definition

912 Area 912
913 Area 913
914 Area 914
915+934 Area $915+$ Area 934

## A. 2 Preparation of the catch information

The preparation of catch data for each of the CRA 4 multi-area sub-stocks was hampered by inconsistent data availability, depending on the period (Table A.2). Reporting from the modern rock lobster statistical areas was not available until 1979, but the fishery was well developed by then, with significant catches taken from at least 1945. Historical catches were available by sub-area from 19631973, but the sub-area definitions used by Annala \& King (1983) (Figure A.1) differed substantially from the rock lobster statistical area definitions used from 1979 onward (Figure A.2). There were also two periods (1945-1962 and 1974-1978) where sub-area catches were not available (Table A.2).

Table A.2: Data availability by time period for CRA 4 multi-area stock assessment. FSU: fisheries statistics unit; QMR: quota management reports; MHR: monthly harvest returns; CELR: catch/effort landing returns. See Bentley et al. (2005) for more information on the historical data sources by year.

| Period | CRA 4 (QMA) data availability | CRA 4 (sub-area) data availability |
| :--- | :--- | :--- |
| 1945-1962 | estimates based on port of landing | not available |
| $1963-1973$ | Annala \& King (1983) | Annala \& King (1983) area definitions |
| $1974-1978$ | estimates based on Annala \& King <br>  <br> (1983) |  |
| not available |  |  |
| $1979-1985$ | FSU estimates | FSU area definitions |
| $1986-1988$ | QMR estimates | FSU area definitions |
| $1989-2000$ | QMR estimates | CELR (FSSU) area definitions |
| 2001-2015 | MHR estimates | CELR (=FSU) area definitions |

An algorithm (Section A.2.1) was developed to convert the Annala \& King (1983) sub-areas into approximate FSU/CELR statistical area definitions, using the statistical area mapping definitions described in Table A. 3 and averaging the observed distributions of catch by sub-area over the first decade of reliable reporting.

Table A.3: Selected mapping of Annala/King statistical areas into the NZ rock lobster statistical areas. CRA 4 statistical areas are marked with *.

| Rock lobster |  | Annala/King statistical area |  |
| :--- | ---: | :---: | :---: |
| statistical area | Area 6 | Area 7 | Area 8 |
| $911 *$ | x |  |  |
| $912^{*}$ | x |  |  |
| $913^{*}$ |  | X |  |
| $914^{*}$ |  | x |  |
| $915^{*}$ |  |  | X |
| 916 |  | x |  |
| 933 |  |  | x |
| $934^{*}$ |  |  | X |



Figure A.1: Map showing Annala \& King statistical areas (from Annala \& King (1983).


Figure A.2: Cropped map of NZ rock lobster statistical areas showing approximate location of the Annala/King statistical areas (labelled in red with red boundary lines).

## A.2.1 Algorithm used to estimate catches in rock lobster statistical areas BEFORE 1979:

The following algorithm was followed to apportion catches to the defined CRA 4 multi-area substocks, based on distributions available when reporting was more accurate:

1. Starting from FSU (in place for rock lobster from January 1979 to June 1988) data set, sum reported catches by statistical area from 1979-80 to year Y (initially set at 1988-89):
$C^{a}=\sum_{y=1979}^{Y} c_{y}^{a} \quad$ where $a$ is one of the rock lobster statistical areas in Table A. 3
2. Calculate the relative proportion of each contributing rock lobster statistical areas among the three Annala/King groupings identified in Table A.3:
$P^{a}=C^{a} / \sum_{a=1}^{N^{A}} C_{a} \quad$ where $N^{A}$ is the number of rock lobster statistical areas in Annala/King area
A. See Table A. 1 for these calculations.
3. By year, calculate the contribution of the five CRA 4 statistical areas based on the proportions estimated in step 2:
$k_{y}^{a}=P^{a} * K_{y}^{A} \quad$ where $K_{y}^{A}$ is the Annala/King catch in area A in year $y(y=1963-1973)$.
Note: only used method codes 8 or 19 in the Annala/King data (these are the potting method)
4. Sum the five contributing CRA 4 statistical areas to get a total CRA 4 estimated catch in each year and divide to get the relative proportion of each CRA 4 statistical area in year $y$ :
$p k_{y}^{a}=k_{y}^{a} / \sum_{a=1}^{5} k_{y}^{a} \quad$ where $a$ is $912,913,914,915$ or 934 in year $y(y=1963-1973)$. See columns
2 to 5 in Table A. 2 for these proportions. Column 6 in Table A. 2 is the sum of catches from the 5 rock lobster statistical areas calculated in this manner.
5. Multiply the resulting proportions by the historical CRA 4 catch in year $y$ to get the scaled statistical area catch from 1963 to 1973:
$c_{y}^{a}=p k_{y}^{a} * C_{y}^{\mathrm{CRA4}}$ where $C_{y}^{\mathrm{CRA4}}$ is the catch value used in the combined CRA 4 stock assessment.

This step is required because the summed $k_{y}^{a}$ is greater than the values stored in CRACE (Bentley et al. 2005) for CRA 4 (which have been used by default in every rock lobster stock assessment; compare columns 6 and 7 in Table A.2).
6. Values for $p k_{y}^{a}$ from 1974 to 1978 were obtained by interpolating between $p k_{1973}^{a}$ and $p k_{1979}^{a}$ (see Table A.3) and then applying the equation in step 5 to the appropriate value of $C_{y}^{\text {CRA4 }}$.
7. Values for $p k_{y}^{a}$ before 1963 were obtained by averaging $\sum_{y=1963}^{y=1973} p k_{y}^{a} / 11$ (see final row in Table A.2) and then applying the equation in step 5 to the appropriate value of $C_{y}^{\mathrm{CRA4}}$.
8. The $p k_{y}^{a}$ proportions derived in Steps 4,6 and 7 were used to allocate the CRA 4 noncommercial catches by statistical area and year in the years before 1979. Proportions based on the reported commercial catches by fishing year from the FSU and CELR systems were used to allocate the non-commercial catches to statistical areas from 1979 onward. The commercial catch distributions were based on the B4_L algorithm (see Starr 2016).

The time series of estimated catches by category of capture are presented by CRA 4 sub-stock in Figure A. 3 and in Table A. 4 (Area 912), Table A. 5 (Area 913), Table A. 6 (Area 914) and Table A. 7 (Area $915+934$ ). Note that the procedure estimates relatively higher catches in the south coast near Wellington (Area $915+934$ ) early in the time period relative to the catches in the other three stocks. CPUE trajectories differed among the four sub-stocks, with Area 914 and Area 915+934 not showing the strong CPUE peak in the late 1990s that was observed in the more northerly sub-stocks (Figure A.4). On the other hand, Area 912 does not show the CPUE peak around 2010 that is present in the other more southerly sub-stocks.

All four sub-stocks show a similar distribution of AW commercial catch (Figure A.5), and these seasonal distributions are used to allocate the commercial and illegal catches to the AW and SS seasons for each of the sub-stocks. Arbitrary seasonal distributions of 0.10 (AW) and 0.90 (SS), consistent with the assumptions used in all rock lobster stock assessments, allocate recreational and customary catches to a season. These seasonal proportions are then used to allocate catches from 1979 onwards into two categories: SL (size limited: the sum of commercial and recreational catches) and NSL (not size limited: the sum of illegal and customary catches) (Figure A.6, Table A. 4 to Table A.7). These are the catch distributions that are used to model each CRA 4 sub-stock.

Table A.1: Percent distribution between adjacent rock lobster statistical areas in the Annala/King area groupings identified in Table A.3.

| Fishing | Area 6 |  | Area 7 |  |  |  | Area 8 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | 911 | 912 | 913 | 914 | 915 | 916 | 933 | 934 |
| 1979 | 60.7 | 39.3 | 44.2 | 55.8 | 25.4 | 53.5 | 20.9 | 0.2 |
| 1980 | 54.6 | 45.4 | 39.3 | 60.7 | 25.5 | 50.7 | 23.4 | 0.3 |
| 1981 | 46.0 | 54.0 | 43.5 | 56.5 | 29.1 | 41.7 | 29.2 | 0.0 |
| 1982 | 51.1 | 48.9 | 36.7 | 63.3 | 34.6 | 34.2 | 31.0 | 0.2 |
| 1983 | 59.3 | 40.8 | 41.0 | 59.0 | 34.3 | 34.2 | 31.1 | 0.4 |
| 1984 | 57.7 | 42.3 | 41.0 | 59.0 | 29.3 | 41.0 | 28.6 | 1.2 |
| 1985 | 55.0 | 45.0 | 36.7 | 63.3 | 35.9 | 40.5 | 22.5 | 1.1 |
| 1986 | 49.5 | 50.5 | 44.0 | 56.0 | 41.0 | 38.7 | 19.4 | 0.9 |
| 1987 | 39.2 | 60.8 | 36.1 | 63.9 | 40.2 | 42.1 | 17.6 | 0.0 |
| 1988 | 47.4 | 52.6 | 37.3 | 62.7 | 42.5 | 38.0 | 19.6 | 0.0 |
| Mean | 52.6 | 47.4 | 39.7 | 60.3 | 33.6 | 40.9 | 25.0 | 0.5 |

Table A.2: Proportions by CRA 4 sub-stock (statistical area) as calculated in Step 4 (above). The final two columns are the sums of annual catch defined in Steps 4 and 5 (above).

| Calendar year | Stock/statistical area |  |  |  | $\sum_{a=1}^{5} k_{y}^{a}$ | $C_{y}^{\text {CRA } 4}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 912 | 913 | 914 | 915+934 |  |  |
| 1963 | 0.265 | 0.124 | 0.189 | 0.422 | 500.2 | 310.3 |
| 1964 | 0.259 | 0.177 | 0.269 | 0.295 | 757.1 | 459.9 |
| 1965 | 0.297 | 0.196 | 0.297 | 0.211 | 938.1 | 581.4 |
| 1966 | 0.195 | 0.225 | 0.341 | 0.239 | 1162.3 | 663.5 |
| 1967 | 0.197 | 0.210 | 0.318 | 0.274 | 892.2 | 512.6 |
| 1968 | 0.195 | 0.219 | 0.332 | 0.255 | 891.3 | 509.6 |
| 1969 | 0.281 | 0.173 | 0.262 | 0.284 | 982.8 | 606.7 |
| 1970 | 0.204 | 0.222 | 0.337 | 0.238 | 972.2 | 559.0 |
| 1971 | 0.159 | 0.222 | 0.337 | 0.281 | 754.1 | 419.3 |
| 1972 | 0.168 | 0.229 | 0.347 | 0.255 | 763.3 | 426.3 |
| 1973 | 0.186 | 0.236 | 0.359 | 0.219 | 661.9 | 373.8 |
| Mean | 0.219 | 0.203 | 0.308 | 0.270 | - |  |

Table A.3: Interpolated proportions by CRA 4 sub-stock (statistical area) as described in Step 6 (above).

| Calendar |  | Stock/statistical area |  |  |
| :--- | ---: | ---: | ---: | ---: |
| Year | $\mathbf{9 1 2}$ | $\mathbf{9 1 3}$ | $\mathbf{9 1 4}$ | $\mathbf{9 1 5}+\mathbf{9 3 4}$ |
| 1973 | 0.186 | 0.236 | 0.359 | 0.219 |
| 1974 | 0.191 | 0.247 | 0.362 | 0.199 |
| 1975 | 0.195 | 0.258 | 0.366 | 0.180 |
| 1976 | 0.200 | 0.269 | 0.370 | 0.160 |
| 1977 | 0.205 | 0.280 | 0.374 | 0.141 |
| 1978 | 0.209 | 0.291 | 0.378 | 0.122 |
| 1979 | 0.214 | 0.302 | 0.382 | 0.102 |



Calendar Year or Fishing Year
-Total - - Commercial ---iliegal


Calendar Year or Fishing Year



Calendar Year or Fishing Year

$$
\begin{array}{ll}
\text { —Total } & -- \text { Commercial } \\
\cdots \text { Customary } & -- \text { Recreational }
\end{array}
$$



Calendar Year or Fishing Year


Figure A.3: Commercial, recreational, illegal and customary catch trajectories for four CRA 4 multi-area sub-stocks defined by rock lobster statistical areas (Table A.1). Year codes are annual before 1978 and from 1979 onward refer to the first year in the statutory 1 April- 31 March fishing year. Catches from January-March 1979 have been added to 1978.


Figure A.4: Commercial catch and annual CPUE for four CRA 4 multi-area sub-stocks defined by rock lobster statistical areas (Table A.1).


Figure A.5: Seasonal proportion of the commercial AW catch by fishing year for four CRA 4 multi-area sub-stocks defined by rock lobster statistical areas (Table A.1). These proportions have been derived from reported landings by month from the FSU or CELR catch reporting systems using the F2_LFX algorithm (Starr 2016).


Figure A.6: The seasonal SL (size-limited) and NSL (non-size-limited) catches (t) for four CRA 4 substocks, defined by rock lobster statistical areas (Table A.1), plotted by fishing year, beginning in 1979.

Table A.4: Estimated catches ( $t$ ) (commercial, recreational including S.111, illegal and customary) for the Area 912 CRA 4 sub-stock, provided annually before 1979 and seasonally (AW and SS) from 1979 to 2015.

| Calendar Year | Commercial Annual | Recrea -tional Annual |  | Illegal Annual | Fishing Year | Commercial |  | Recreational |  | Customary |  | Illegal |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |  | AW | SS | AW | SS | AW | SS | AW | SS |
| 1945 | 55.7 | 2.6 | 4.4 | 10.0 | 1979 | 31.3 | 76.3 | 0.8 | 7.1 | 0.4 | 3.8 | 2.7 | 6.6 |
| 1946 | 49.3 | 2.8 | 4.4 | 8.8 | 1980 | 68.7 | 128.5 | 1.1 | 10.0 | 0.6 | 5.8 | 7.8 | 14.7 |
| 1947 | 55.5 | 3.0 | 4.4 | 9.9 | 1981 | 79.2 | 139.7 | 1.3 | 11.8 | 0.7 | 6.4 | 14.2 | 25.0 |
| 1948 | 55.4 | 3.1 | 4.4 | 9.9 | 1982 | 78.5 | 141.1 | 1.0 | 9.2 | 0.5 | 4.6 | 14.1 | 25.2 |
| 1949 | 59.9 | 3.3 | 4.4 | 10.7 | 1983 | 68.5 | 117.4 | 0.7 | 6.3 | 0.4 | 3.6 | 12.3 | 21.0 |
| 1950 | 110.2 | 3.4 | 4.4 | 19.7 | 1984 | 71.0 | 145.5 | 0.8 | 7.5 | 0.5 | 4.5 | 12.7 | 26.0 |
| 1951 | 147.4 | 3.6 | 4.4 | 26.4 | 1985 | 64.5 | 164.4 | 1.0 | 8.6 | 0.5 | 4.9 | 11.5 | 29.4 |
| 1952 | 143.0 | 3.8 | 4.4 | 25.6 | 1986 | 50.5 | 156.8 | 0.8 | 7.3 | 0.4 | 3.9 | 9.0 | 28.1 |
| 1953 | 148.5 | 3.9 | 4.4 | 26.6 | 1987 | 56.3 | 122.9 | 0.7 | 5.9 | 0.4 | 3.5 | 10.1 | 22.0 |
| 1954 | 145.8 | 4.1 | 4.4 | 26.1 | 1988 | 40.8 | 93.9 | 0.5 | 4.6 | 0.4 | 3.2 | 7.3 | 16.8 |
| 1955 | 110.2 | 4.2 | 4.4 | 19.7 | 1989 | 45.5 | 129.0 | 0.7 | 6.2 | 0.5 | 4.1 | 8.1 | 23.1 |
| 1956 | 94.9 | 4.4 | 4.4 | 17.0 | 1990 | 52.5 | 95.4 | 0.8 | 7.0 | 0.6 | 5.1 | 16.1 | 29.2 |
| 1957 | 71.7 | 4.5 | 4.4 | 12.8 | 1991 | 52.8 | 114.7 | 0.9 | 7.8 | 0.6 | 5.7 | 9.5 | 20.6 |
| 1958 | 74.5 | 4.7 | 4.4 | 13.3 | 1992 | 51.5 | 97.8 | 0.8 | 7.3 | 0.6 | 5.4 | 3.1 | 5.9 |
| 1959 | 64.3 | 4.9 | 4.4 | 11.5 | 1993 | 43.4 | 73.5 | 0.7 | 6.4 | 0.5 | 4.3 | 4.4 | 7.5 |
| 1960 | 79.2 | 5.0 | 4.4 | 14.2 | 1994 | 53.5 | 54.0 | 0.8 | 7.0 | 0.4 | 3.9 | 7.6 | 7.7 |
| 1961 | 91.8 | 5.2 | 4.4 | 16.4 | 1995 | 76.9 | 34.5 | 1.1 | 9.8 | 0.5 | 4.1 | 10.1 | 4.5 |
| 1962 | 109.7 | 5.3 | 4.4 | 19.6 | 1996 | 105.7 | 15.6 | 1.5 | 13.9 | 0.5 | 4.4 | 16.1 | 2.4 |
| 1963 | 82.2 | 6.7 | 5.3 | 14.7 | 1997 | 122.6 | 2.6 | 1.7 | 15.2 | 0.5 | 4.6 | 18.1 | 0.4 |
| 1964 | 119.2 | 6.7 | 5.2 | 21.3 | 1998 | 140.4 | 14.1 | 2.7 | 24.1 | 0.6 | 5.6 | 19.8 | 2.0 |
| 1965 | 172.6 | 7.9 | 5.9 | 30.9 | 1999 | 147.1 | 5.9 | 1.9 | 17.3 | 0.5 | 4.8 | 17.0 | 0.7 |
| 1966 | 129.6 | 5.3 | 3.9 | 23.2 | 2000 | 131.1 | 23.4 | 2.2 | 19.6 | 0.5 | 4.8 | 14.6 | 2.6 |
| 1967 | 101.1 | 5.5 | 3.9 | 18.1 | 2001 | 92.8 | 34.5 | 1.4 | 12.6 | 0.4 | 4.0 | 10.0 | 3.7 |
| 1968 | 99.1 | 5.6 | 3.9 | 17.7 | 2002 | 104.6 | 30.2 | 1.6 | 14.1 | 0.5 | 4.2 | 10.9 | 3.1 |
| 1969 | 170.3 | 8.3 | 5.6 | 30.5 | 2003 | 63.0 | 47.9 | 1.2 | 11.2 | 0.4 | 3.5 | 5.5 | 4.2 |
| 1970 | 113.8 | 6.2 | 4.1 | 20.4 | 2004 | 33.4 | 55.3 | 0.8 | 7.2 | 0.3 | 2.8 | 2.3 | 3.9 |
| 1971 | 66.9 | 4.9 | 3.2 | 12.0 | 2005 | 13.6 | 35.3 | 0.4 | 3.3 | 0.2 | 1.7 | 1.1 | 2.8 |
| 1972 | 71.8 | 5.3 | 3.4 | 12.8 | 2006 | 11.0 | 42.6 | 0.4 | 3.6 | 0.2 | 2.2 | 1.0 | 3.8 |
| 1973 | 69.7 | 6.1 | 3.7 | 12.5 | 2007 | 9.4 | 40.8 | 0.5 | 4.3 | 0.3 | 2.9 | 1.2 | 5.2 |
| 1974 | 71.6 | 6.3 | 3.8 | 9.2 | 2008 | 12.0 | 34.9 | 0.7 | 6.4 | 0.4 | 3.4 | 1.9 | 5.6 |
| 1975 | 79.0 | 6.6 | 3.9 | 19.1 | 2009 | 11.9 | 33.0 | 0.9 | 7.7 | 0.3 | 3.1 | 1.8 | 5.0 |
| 1976 | 91.2 | 6.9 | 4.0 | 17.7 | 2010 | 28.0 | 30.7 | 0.7 | 6.5 | 0.3 | 2.5 | 2.7 | 3.0 |
| 1977 | 89.6 | 7.2 | 4.1 | 23.0 | 2011 | 18.6 | 26.6 | 0.6 | 5.1 | 0.2 | 1.7 | 1.6 | 2.3 |
| 1978 | 103.8 | 7.6 | 4.2 | 26.6 | 2012 | 17.4 | 29.0 | 0.6 | 5.6 | 0.2 | 1.8 | 1.5 | 2.5 |
|  |  |  |  |  | 2013 | 13.7 | 29.6 | 0.5 | 4.2 | 0.2 | 1.6 | 1.1 | 2.4 |
|  |  |  |  |  | 2014 | 8.3 | 25.9 | 0.3 | 3.1 | 0.1 | 1.3 | 0.7 | 2.2 |
|  |  |  |  |  | 2015 | 7.6 | 43.2 | 0.4 | 3.9 | 0.2 | 2.1 | 0.7 | 3.9 |

Table A.5: Estimated catches ( $\mathbf{t}$ ) (commercial, recreational including S.111, illegal and customary) for the Area 913 CRA 4 sub-stock, provided annually before 1979 and seasonally (AW and SS) from 1979 to 2015.

| Calendar Year | Commercial Annual | Recrea -tional Annual |  | Illegal Annual | Fishing Year | Commercial |  | Recreational |  | Customary |  | Illegal |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |  | AW | SS | AW | SS | AW | SS | AW | SS |
| 1945 | 51.7 | 2.4 | 4.1 | 9.3 | 1979 | 58.0 | 94.3 | 1.1 | 10.0 | 0.6 | 5.4 | 5.0 | 8.2 |
| 1946 | 45.8 | 2.6 | 4.1 | 8.2 | 1980 | 62.9 | 68.7 | 0.7 | 6.7 | 0.4 | 3.9 | 7.2 | 7.8 |
| 1947 | 51.5 | 2.7 | 4.1 | 9.2 | 1981 | 67.5 | 71.4 | 0.8 | 7.5 | 0.5 | 4.1 | 12.1 | 12.8 |
| 1948 | 51.4 | 2.9 | 4.1 | 9.2 | 1982 | 69.2 | 117.0 | 0.9 | 7.8 | 0.4 | 3.9 | 12.4 | 20.9 |
| 1949 | 55.6 | 3.0 | 4.1 | 10.0 | 1983 | 109.3 | 152.4 | 1.0 | 8.9 | 0.6 | 5.0 | 19.6 | 27.3 |
| 1950 | 102.2 | 3.2 | 4.1 | 18.3 | 1984 | 85.4 | 136.7 | 0.9 | 7.7 | 0.5 | 4.6 | 15.3 | 24.5 |
| 1951 | 136.8 | 3.3 | 4.1 | 24.5 | 1985 | 57.4 | 122.7 | 0.8 | 6.8 | 0.4 | 3.8 | 10.3 | 22.0 |
| 1952 | 132.8 | 3.5 | 4.1 | 23.8 | 1986 | 75.1 | 202.7 | 1.1 | 9.8 | 0.6 | 5.3 | 13.4 | 36.3 |
| 1953 | 137.8 | 3.6 | 4.1 | 24.7 | 1987 | 78.8 | 153.7 | 0.8 | 7.6 | 0.5 | 4.5 | 14.1 | 27.5 |
| 1954 | 135.4 | 3.8 | 4.1 | 24.2 | 1988 | 64.6 | 142.0 | 0.8 | 7.1 | 0.5 | 4.9 | 11.6 | 25.4 |
| 1955 | 102.3 | 3.9 | 4.1 | 18.3 | 1989 | 74.1 | 193.3 | 1.0 | 9.4 | 0.7 | 6.3 | 13.3 | 34.6 |
| 1956 | 88.1 | 4.1 | 4.1 | 15.8 | 1990 | 46.6 | 107.6 | 0.8 | 7.3 | 0.6 | 5.3 | 14.3 | 32.9 |
| 1957 | 66.6 | 4.2 | 4.1 | 11.9 | 1991 | 52.5 | 102.8 | 0.8 | 7.3 | 0.6 | 5.3 | 9.4 | 18.4 |
| 1958 | 69.2 | 4.4 | 4.1 | 12.4 | 1992 | 43.9 | 86.5 | 0.7 | 6.4 | 0.5 | 4.7 | 2.7 | 5.2 |
| 1959 | 59.7 | 4.5 | 4.1 | 10.7 | 1993 | 64.5 | 77.0 | 0.9 | 7.7 | 0.6 | 5.2 | 6.6 | 7.8 |
| 1960 | 73.5 | 4.7 | 4.1 | 13.2 | 1994 | 49.4 | 70.8 | 0.9 | 7.9 | 0.5 | 4.4 | 7.1 | 10.1 |
| 1961 | 85.3 | 4.8 | 4.1 | 15.3 | 1995 | 60.9 | 51.6 | 1.1 | 9.9 | 0.5 | 4.2 | 8.0 | 6.8 |
| 1962 | 101.8 | 5.0 | 4.1 | 18.2 | 1996 | 93.3 | 3.4 | 1.2 | 11.0 | 0.4 | 3.5 | 14.2 | 0.5 |
| 1963 | 38.6 | 3.1 | 2.5 | 6.9 | 1997 | 99.3 | 8.4 | 1.5 | 13.1 | 0.4 | 4.0 | 14.6 | 1.2 |
| 1964 | 81.4 | 4.6 | 3.5 | 14.6 | 1998 | 107.3 | 0.9 | 1.9 | 16.9 | 0.4 | 3.9 | 15.1 | 0.1 |
| 1965 | 113.8 | 5.2 | 3.9 | 20.4 | 1999 | 126.9 | 2.2 | 1.6 | 14.6 | 0.4 | 4.0 | 14.7 | 0.3 |
| 1966 | 149.2 | 6.2 | 4.5 | 26.7 | 2000 | 116.6 | 18.0 | 1.9 | 17.0 | 0.5 | 4.2 | 13.0 | 2.0 |
| 1967 | 107.6 | 5.9 | 4.2 | 19.3 | 2001 | 97.5 | 26.2 | 1.4 | 12.3 | 0.4 | 3.9 | 10.5 | 2.8 |
| 1968 | 111.5 | 6.3 | 4.4 | 20.0 | 2002 | 116.8 | 38.8 | 1.8 | 16.3 | 0.5 | 4.9 | 12.2 | 4.0 |
| 1969 | 105.0 | 5.1 | 3.5 | 18.8 | 2003 | 111.1 | 72.8 | 2.1 | 18.6 | 0.6 | 5.7 | 9.7 | 6.3 |
| 1970 | 124.1 | 6.7 | 4.4 | 22.2 | 2004 | 75.4 | 86.7 | 1.5 | 13.1 | 0.6 | 5.1 | 5.3 | 6.1 |
| 1971 | 93.2 | 6.9 | 4.4 | 16.7 | 2005 | 46.6 | 59.9 | 0.8 | 7.1 | 0.4 | 3.8 | 3.7 | 4.8 |
| 1972 | 97.7 | 7.3 | 4.6 | 17.5 | 2006 | 28.7 | 74.7 | 0.8 | 7.0 | 0.5 | 4.2 | 2.6 | 6.7 |
| 1973 | 88.4 | 7.7 | 4.7 | 15.8 | 2007 | 12.7 | 53.4 | 0.6 | 5.7 | 0.4 | 3.8 | 1.6 | 6.8 |
| 1974 | 92.8 | 8.2 | 4.9 | 11.9 | 2008 | 15.1 | 56.8 | 1.1 | 9.8 | 0.6 | 5.2 | 2.4 | 9.1 |
| 1975 | 104.4 | 8.8 | 5.2 | 25.3 | 2009 | 31.3 | 36.4 | 1.3 | 11.7 | 0.5 | 4.6 | 4.8 | 5.6 |
| 1976 | 122.8 | 9.3 | 5.4 | 23.9 | 2010 | 45.1 | 46.1 | 1.1 | 10.1 | 0.4 | 4.0 | 4.4 | 4.4 |
| 1977 | 122.8 | 9.9 | 5.6 | 31.5 | 2011 | 83.9 | 53.1 | 1.7 | 15.6 | 0.6 | 5.3 | 7.2 | 4.6 |
| 1978 | 144.6 | 10.5 | 5.8 | 37.1 | 2012 | 76.4 | 72.0 | 2.0 | 18.1 | 0.6 | 5.7 | 6.6 | 6.2 |
|  |  |  |  |  | 2013 | 60.0 | 76.2 | 1.5 | 13.3 | 0.5 | 4.9 | 4.8 | 6.1 |
|  |  |  |  |  | 2014 | 29.3 | 46.2 | 0.8 | 6.8 | 0.3 | 2.9 | 2.5 | 4.0 |
|  |  |  |  |  | 2015 | 20.7 | 81.1 | 0.9 | 7.8 | 0.5 | 4.2 | 1.9 | 7.4 |

Table A.6: Estimated catches ( $t$ ) (commercial, recreational including S.111, illegal and customary) for the Area 914 CRA 4 sub-stock, provided annually before 1979 and seasonally (AW and SS) from 1979 to 2015.

| Calendar | Commercial | Recrea -tional | Customary | Illegal | Fishing | Commercial |  | Recreational |  | Customary |  | Illegal |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | Annual | Annual | Annual | Annual | Year | AW | SS | AW | SS | AW | SS | AW | SS |
| 1945 | 78.5 | 3.7 | 6.2 | 14.0 | 1979 | 58.7 | 133.6 | 1.4 | 12.7 | 0.8 | 6.9 | 5.1 | 11.6 |
| 1946 | 69.4 | 3.9 | 6.2 | 12.4 | 1980 | 78.3 | 125.4 | 1.2 | 10.4 | 0.7 | 6.0 | 8.9 | 14.3 |
| 1947 | 78.1 | 4.2 | 6.2 | 14.0 | 1981 | 62.4 | 117.7 | 1.1 | 9.7 | 0.6 | 5.3 | 11.2 | 21.1 |
| 1948 | 78.0 | 4.4 | 6.2 | 14.0 | 1982 | 118.6 | 202.5 | 1.5 | 13.4 | 0.8 | 6.8 | 21.2 | 36.2 |
| 1949 | 84.4 | 4.6 | 6.2 | 15.1 | 1983 | 153.8 | 222.7 | 1.4 | 12.8 | 0.8 | 7.2 | 27.5 | 39.9 |
| 1950 | 155.1 | 4.8 | 6.2 | 27.8 | 1984 | 141.6 | 178.3 | 1.2 | 11.1 | 0.7 | 6.7 | 25.4 | 31.9 |
| 1951 | 207.5 | 5.1 | 6.2 | 37.1 | 1985 | 103.6 | 207.4 | 1.3 | 11.7 | 0.7 | 6.6 | 18.5 | 37.1 |
| 1952 | 201.4 | 5.3 | 6.2 | 36.0 | 1986 | 112.3 | 241.7 | 1.4 | 12.4 | 0.7 | 6.7 | 20.1 | 43.3 |
| 1953 | 209.1 | 5.5 | 6.2 | 37.4 | 1987 | 108.5 | 302.8 | 1.5 | 13.5 | 0.9 | 8.0 | 19.4 | 54.2 |
| 1954 | 205.3 | 5.7 | 6.2 | 36.7 | 1988 | 116.4 | 231.5 | 1.3 | 11.9 | 0.9 | 8.2 | 20.8 | 41.4 |
| 1955 | 155.2 | 6.0 | 6.2 | 27.8 | 1989 | 98.7 | 157.6 | 1.0 | 9.0 | 0.7 | 6.1 | 17.7 | 28.2 |
| 1956 | 133.7 | 6.2 | 6.2 | 23.9 | 1990 | 43.8 | 121.9 | 0.9 | 7.8 | 0.6 | 5.7 | 13.4 | 37.3 |
| 1957 | 100.9 | 6.4 | 6.2 | 18.1 | 1991 | 48.7 | 110.7 | 0.8 | 7.5 | 0.6 | 5.4 | 8.7 | 19.8 |
| 1958 | 104.9 | 6.6 | 6.2 | 18.8 | 1992 | 65.3 | 96.3 | 0.9 | 7.9 | 0.7 | 5.9 | 4.0 | 5.8 |
| 1959 | 90.6 | 6.9 | 6.2 | 16.2 | 1993 | 99.9 | 80.7 | 1.1 | 9.8 | 0.7 | 6.6 | 10.2 | 8.2 |
| 1960 | 111.5 | 7.1 | 6.2 | 19.9 | 1994 | 136.5 | 68.1 | 1.5 | 13.4 | 0.8 | 7.5 | 19.5 | 9.7 |
| 1961 | 129.3 | 7.3 | 6.2 | 23.1 | 1995 | 178.5 | 49.8 | 2.2 | 20.1 | 0.9 | 8.4 | 23.4 | 6.5 |
| 1962 | 154.4 | 7.5 | 6.2 | 27.6 | 1996 | 214.7 | 12.5 | 2.9 | 26.0 | 0.9 | 8.3 | 32.6 | 1.9 |
| 1963 | 58.5 | 4.7 | 3.8 | 10.5 | 1997 | 209.4 | 11.2 | 3.0 | 26.8 | 0.9 | 8.1 | 30.9 | 1.7 |
| 1964 | 123.5 | 7.0 | 5.4 | 22.1 | 1998 | 170.1 | 18.4 | 3.3 | 29.4 | 0.8 | 6.9 | 24.0 | 2.6 |
| 1965 | 172.6 | 7.9 | 5.9 | 30.9 | 1999 | 206.9 | 21.8 | 2.9 | 25.8 | 0.8 | 7.1 | 24.0 | 2.5 |
| 1966 | 226.4 | 9.3 | 6.8 | 40.5 | 2000 | 194.1 | 22.7 | 3.0 | 27.4 | 0.8 | 6.8 | 21.7 | 2.5 |
| 1967 | 163.3 | 8.9 | 6.4 | 29.2 | 2001 | 205.7 | 36.9 | 2.7 | 24.1 | 0.8 | 7.6 | 22.2 | 4.0 |
| 1968 | 169.2 | 9.6 | 6.6 | 30.3 | 2002 | 154.0 | 56.1 | 2.4 | 22.0 | 0.7 | 6.6 | 16.0 | 5.8 |
| 1969 | 159.2 | 7.8 | 5.2 | 28.5 | 2003 | 156.0 | 78.8 | 2.6 | 23.7 | 0.8 | 7.3 | 13.5 | 6.8 |
| 1970 | 188.2 | 10.2 | 6.7 | 33.7 | 2004 | 128.1 | 149.9 | 2.5 | 22.4 | 1.0 | 8.8 | 9.0 | 10.5 |
| 1971 | 141.4 | 10.5 | 6.7 | 25.3 | 2005 | 113.4 | 163.9 | 2.1 | 18.5 | 1.1 | 9.9 | 9.0 | 13.0 |
| 1972 | 148.1 | 11.0 | 6.9 | 26.5 | 2006 | 46.4 | 148.9 | 1.5 | 13.1 | 0.9 | 7.9 | 4.2 | 13.4 |
| 1973 | 134.0 | 11.6 | 7.2 | 24.0 | 2007 | 29.3 | 91.8 | 1.2 | 10.5 | 0.8 | 6.9 | 3.7 | 11.7 |
| 1974 | 135.9 | 12.0 | 7.2 | 17.4 | 2008 | 28.2 | 60.7 | 1.3 | 12.1 | 0.7 | 6.4 | 4.5 | 9.7 |
| 1975 | 148.0 | 12.4 | 7.3 | 35.8 | 2009 | 41.1 | 46.4 | 1.7 | 15.1 | 0.7 | 6.0 | 6.3 | 7.1 |
| 1976 | 168.8 | 12.8 | 7.4 | 32.8 | 2010 | 88.6 | 99.3 | 2.3 | 20.8 | 0.9 | 8.2 | 8.5 | 9.6 |
| 1977 | 163.8 | 13.2 | 7.5 | 42.1 | 2011 | 131.5 | 97.4 | 2.9 | 26.1 | 1.0 | 8.8 | 11.3 | 8.4 |
| 1978 | 187.6 | 13.7 | 7.6 | 48.1 | 2012 | 115.5 | 109.6 | 3.0 | 27.4 | 1.0 | 8.7 | 9.9 | 9.4 |
|  |  |  |  |  | 2013 | 132.1 | 154.2 | 3.1 | 28.1 | 1.1 | 10.3 | 10.6 | 12.4 |
|  |  |  |  |  | 2014 | 106.3 | 166.2 | 2.7 | 24.5 | 1.2 | 10.5 | 9.1 | 14.3 |
|  |  |  |  |  | 2015 | 53.8 | 150.7 | 1.7 | 15.7 | 0.9 | 8.4 | 4.9 | 13.8 |

Table A.7: Estimated catches (t) (commercial, recreational including S.111, illegal and customary) for the Area 915+934 CRA 4 sub-stock, provided annually before 1979 and seasonally (AW and SS) from 1979 to 2015.

| Calendar | Commercial | Recrea -tional | Customary | Illegal | Fishing | Commercial |  | Recreational |  | Customary |  | Illegal |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | Annual | Annual | Annual | Annual | Year | AW | SS | AW | SS | AW | SS | AW | SS |
| 1945 | 68.8 | 3.3 | 5.4 | 12.3 | 1979 | 11.9 | 39.5 | 0.4 | 3.4 | 0.2 | 1.8 | 1.0 | 3.4 |
| 1946 | 60.9 | 3.4 | 5.4 | 10.9 | 1980 | 14.3 | 61.0 | 0.4 | 3.8 | 0.2 | 2.2 | 1.6 | 7.0 |
| 1947 | 68.5 | 3.6 | 5.4 | 12.3 | 1981 | 20.0 | 56.4 | 0.5 | 4.1 | 0.2 | 2.2 | 3.6 | 10.1 |
| 1948 | 68.4 | 3.8 | 5.4 | 12.2 | 1982 | 40.3 | 86.3 | 0.6 | 5.3 | 0.3 | 2.7 | 7.2 | 15.5 |
| 1949 | 74.0 | 4.0 | 5.4 | 13.2 | 1983 | 40.9 | 75.4 | 0.4 | 3.9 | 0.2 | 2.2 | 7.3 | 13.5 |
| 1950 | 136.0 | 4.2 | 5.4 | 24.3 | 1984 | 43.2 | 61.4 | 0.4 | 3.6 | 0.2 | 2.2 | 7.7 | 11.0 |
| 1951 | 182.0 | 4.4 | 5.4 | 32.6 | 1985 | 46.7 | 81.4 | 0.5 | 4.8 | 0.3 | 2.7 | 8.4 | 14.6 |
| 1952 | 176.6 | 4.6 | 5.4 | 31.6 | 1986 | 32.9 | 75.5 | 0.4 | 3.8 | 0.2 | 2.1 | 5.9 | 13.5 |
| 1953 | 183.4 | 4.8 | 5.4 | 32.8 | 1987 | 34.5 | 71.7 | 0.4 | 3.5 | 0.2 | 2.1 | 6.2 | 12.8 |
| 1954 | 180.1 | 5.0 | 5.4 | 32.2 | 1988 | 20.5 | 55.5 | 0.3 | 2.6 | 0.2 | 1.8 | 3.7 | 9.9 |
| 1955 | 136.1 | 5.2 | 5.4 | 24.4 | 1989 | 22.2 | 38.1 | 0.2 | 2.1 | 0.2 | 1.4 | 4.0 | 6.8 |
| 1956 | 117.2 | 5.4 | 5.4 | 21.0 | 1990 | 18.9 | 36.5 | 0.3 | 2.6 | 0.2 | 1.9 | 5.8 | 11.2 |
| 1957 | 88.5 | 5.6 | 5.4 | 15.8 | 1991 | 18.3 | 30.0 | 0.3 | 2.3 | 0.2 | 1.6 | 3.3 | 5.4 |
| 1958 | 92.0 | 5.8 | 5.4 | 16.5 | 1992 | 21.8 | 32.7 | 0.3 | 2.7 | 0.2 | 2.0 | 1.3 | 2.0 |
| 1959 | 79.4 | 6.0 | 5.4 | 14.2 | 1993 | 25.4 | 27.6 | 0.3 | 2.9 | 0.2 | 1.9 | 2.6 | 2.8 |
| 1960 | 97.8 | 6.2 | 5.4 | 17.5 | 1994 | 32.8 | 25.2 | 0.4 | 3.8 | 0.2 | 2.1 | 4.7 | 3.6 |
| 1961 | 113.4 | 6.4 | 5.4 | 20.3 | 1995 | 27.5 | 7.5 | 0.3 | 3.1 | 0.1 | 1.3 | 3.6 | 1.0 |
| 1962 | 135.5 | 6.6 | 5.4 | 24.2 | 1996 | 43.6 | 4.8 | 0.6 | 5.5 | 0.2 | 1.8 | 6.6 | 0.7 |
| 1963 | 131.0 | 10.6 | 8.4 | 23.5 | 1997 | 33.5 | 3.4 | 0.5 | 4.5 | 0.2 | 1.4 | 4.9 | 0.5 |
| 1964 | 135.7 | 7.6 | 5.9 | 24.3 | 1998 | 31.7 | 10.3 | 0.7 | 6.6 | 0.2 | 1.5 | 4.5 | 1.5 |
| 1965 | 122.4 | 5.6 | 4.2 | 21.9 | 1999 | 63.1 | 2.6 | 0.8 | 7.4 | 0.2 | 2.1 | 7.3 | 0.3 |
| 1966 | 158.3 | 6.5 | 4.8 | 28.3 | 2000 | 54.8 | 13.1 | 1.0 | 8.6 | 0.2 | 2.1 | 6.1 | 1.5 |
| 1967 | 140.6 | 7.7 | 5.5 | 25.2 | 2001 | 74.1 | 6.3 | 0.9 | 8.0 | 0.3 | 2.5 | 8.0 | 0.7 |
| 1968 | 129.7 | 7.3 | 5.1 | 23.2 | 2002 | 54.7 | 20.5 | 0.9 | 7.9 | 0.3 | 2.4 | 5.7 | 2.1 |
| 1969 | 172.3 | 8.4 | 5.7 | 30.8 | 2003 | 26.8 | 19.3 | 0.5 | 4.7 | 0.2 | 1.4 | 2.3 | 1.7 |
| 1970 | 132.9 | 7.2 | 4.8 | 23.8 | 2004 | 14.6 | 26.6 | 0.4 | 3.3 | 0.1 | 1.3 | 1.0 | 1.9 |
| 1971 | 117.8 | 8.7 | 5.6 | 21.1 | 2005 | 36.2 | 35.2 | 0.5 | 4.8 | 0.3 | 2.5 | 2.9 | 2.8 |
| 1972 | 108.8 | 8.1 | 5.1 | 19.5 | 2006 | 23.5 | 68.8 | 0.7 | 6.2 | 0.4 | 3.7 | 2.1 | 6.2 |
| 1973 | 81.8 | 7.1 | 4.4 | 14.6 | 2007 | 25.2 | 52.7 | 0.7 | 6.7 | 0.5 | 4.4 | 3.2 | 6.7 |
| 1974 | 74.7 | 6.6 | 4.0 | 9.6 | 2008 | 17.9 | 23.9 | 0.6 | 5.7 | 0.3 | 3.0 | 2.9 | 3.8 |
| 1975 | 72.7 | 6.1 | 3.6 | 17.6 | 2009 | 33.9 | 28.3 | 1.2 | 10.7 | 0.5 | 4.3 | 5.2 | 4.3 |
| 1976 | 73.1 | 5.6 | 3.2 | 14.2 | 2010 | 37.5 | 39.6 | 0.9 | 8.5 | 0.4 | 3.3 | 3.6 | 3.8 |
| 1977 | 61.7 | 5.0 | 2.8 | 15.9 | 2011 | 22.4 | 32.8 | 0.7 | 6.3 | 0.2 | 2.1 | 1.9 | 2.8 |
| 1978 | 60.3 | 4.4 | 2.4 | 15.5 | 2012 | 16.0 | 30.4 | 0.6 | 5.6 | 0.2 | 1.8 | 1.4 | 2.6 |
|  |  |  |  |  | 2013 | 13.1 | 20.5 | 0.4 | 3.3 | 0.1 | 1.2 | 1.0 | 1.6 |
|  |  |  |  |  | 2014 | 34.5 | 48.8 | 0.8 | 7.5 | 0.4 | 3.2 | 3.0 | 4.2 |
|  |  |  |  |  | 2015 | 27.8 | 53.3 | 0.7 | 6.2 | 0.4 | 3.3 | 2.5 | 4.9 |

## A. 3 Catch Rate Information

## A.3.1 FSU \& CELR CPUE IndICES

Catch and effort data from the FSU and CELR systems were obtained from MPI in September 2016 (Replog 10736), loaded into the CRACE database and processed using standard error checks (Bentley et al. 2005). Data spanned the period from 1 April 1979 through to 31 March 2016.

Data preparation used the F2-LFX procedure (Starr 2016). The F2 algorithm corrects the monthly estimated catch taken by a vessel in a statistical area using a "vessel correction factor" ( $v c f$ : the ratio of landed catch to estimated catch for one vessel in one year) (Starr 2016; Starr et al. 2012), and discards from the analysis those vessels with $v c f$ less than 0.8 or greater than 1.2. The F2-LFX procedure scales the estimated catches to the combined "L" (LFR), "X" (discarded to sea) and "F" (Section 111 recreational catch) destination codes.

The CPUE standardisation procedure used sequential six-month periods as a forced explanatory variable. The only explanatory variable available for the single statistical area analyses is [month] of capture. The variable [statistical_area] was added for the Area 915+934 analysis. These analyses estimate separate relative [month] effects in each half-year period by using, as the reference [month], the [month] in each period with the lowest standard error.

## A.3.1.1 Area 912

The Area 912 data set shows a diminishing number of records (Table A.8). The total deviance explained by the Area 912 model was $36 \%$ (Table A.9), with only month available for standardisation apart from the time period variable. Residual patterns showed some deviation from the lognormal assumption at both tails of the residual distribution (Figure A.7).

The month categorical variable in the CRA 4 seasonal CPUE analysis appears to be cyclical, with a winter peak in June and an early summer peak in November (Figure A.8). Both the Area 912 AW and SS CPUE series showed similar patterns, with the AW series having lower absolute catch rates (Figure A.9, Table A.10). Both series peak in the late 1990s and a second peak around 2010, seen in other stocks, is not well defined here (Figure A.9). The SS series had a larger associated error than the AW series, particularly during the peak in the late 1990s, reflecting the relatively smaller amount of data in the SS series in those years (Figure A.9).

Table A.8. Number of vessel/month records in the dataset used to calculate the Area 912 CRA 4 sub-stock CPUE time series (based on the F2_LFX algorithm). '-': no data.

| Fishing year | AW |  |  |  |  |  |  |  |  |  |  |  | SS |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Apr | May | Jun | Jul | Aug | Sep | Total | Oct | Nov | Dec | Jan | Feb | Mar | Total |
| 1979 | 5 | 3 | 19 | 20 | 21 | 21 | 89 | 23 | 28 | 26 | 23 | 23 | 13 | 136 |
| 1980 | 11 | 6 | 21 | 22 | 21 | 20 | 101 | 27 | 24 | 25 | 26 | 26 | 21 | 149 |
| 1981 | 14 | 8 | 18 | 21 | 23 | 25 | 109 | 25 | 26 | 25 | 27 | 22 | 21 | 146 |
| 1982 | 14 | 13 | 22 | 22 | 25 | 26 | 122 | 25 | 26 | 26 | 25 | 27 | 18 | 147 |
| 1983 | 10 | 13 | 21 | 18 | 23 | 24 | 109 | 24 | 23 | 25 | 24 | 21 | 20 | 137 |
| 1984 | 9 | 11 | 21 | 19 | 19 | 20 | 99 | 21 | 21 | 23 | 23 | 18 | 12 | 118 |
| 1985 | 7 | 13 | 22 | 23 | 23 | 24 | 112 | 24 | 24 | 23 | 23 | 23 | 16 | 133 |
| 1986 | 7 | 12 | 20 | 20 | 20 | 23 | 102 | 23 | 22 | 24 | 23 | 22 | 13 | 127 |
| 1987 | 10 | 16 | 19 | 16 | 17 | 20 | 98 | 21 | 22 | 20 | 21 | 20 | 17 | 121 |
| 1988 | 10 | 12 | 18 | 16 | 19 | 19 | 94 | 18 | 18 | 18 | 17 | 17 | 17 | 105 |
| 1989 | 11 | 13 | 17 | 15 | 19 | 19 | 94 | 20 | 20 | 24 | 18 | 18 | 15 | 115 |
| 1990 | 6 | 11 | 19 | 19 | 18 | 20 | 93 | 22 | 20 | 20 | 19 | 18 | 14 | 113 |
| 1991 | 13 | 17 | 20 | 23 | 22 | 22 | 117 | 23 | 24 | 23 | 21 | 20 | 15 | 126 |
| 1992 | 13 | 22 | 26 | 26 | 25 | 27 | 139 | 29 | 28 | 27 | 22 | 19 | 19 | 144 |
| 1993 | 17 | 26 | 28 | 24 | 21 | 20 | 136 | 25 | 21 | 18 | 13 | 14 | 8 | 99 |
| 1994 | 13 | 14 | 17 | 15 | 16 | 16 | 91 | 3 | 15 | 15 | 7 | 1 | 6 | 47 |
| 1995 | 12 | 13 | 13 | 14 | 14 | 14 | 80 | 12 | 11 | 3 | 2 | 3 | 10 | 41 |
| 1996 | 10 | 11 | 12 | 12 | 15 | 14 | 74 | 7 | 3 | 3 | 1 | 1 | 7 | 22 |
| 1997 | 10 | 11 | 11 | 11 | 11 | 9 | 63 | 2 | - | 2 | 1 | - | 5 | 10 |
| 1998 | 9 | 10 | 13 | 15 | 15 | 16 | 78 | 5 | 2 | 2 | - | 5 | 2 | 16 |
| 1999 | 7 | 13 | 13 | 13 | 14 | 13 | 73 | 4 | 3 | 2 | 1 | 2 | 4 | 16 |
| 2000 | 10 | 14 | 14 | 15 | 15 | 11 | 79 | 10 | 2 | 4 | 2 | 3 | 5 | 26 |
| 2001 | 9 | 11 | 13 | 14 | 13 | 15 | 75 | 10 | 7 | 3 | 2 | 2 | 3 | 27 |
| 2002 | 13 | 13 | 13 | 13 | 12 | 13 | 77 | 8 | 5 | 5 | 6 | 8 | 9 | 41 |
| 2003 | 6 | 11 | 13 | 14 | 10 | 14 | 68 | 16 | 10 | 6 | 8 | 6 | 6 | 52 |
| 2004 | 9 | 7 | 14 | 13 | 11 | 13 | 67 | 14 | 13 | 11 | 8 | 9 | 9 | 64 |
| 2005 | 1 | 7 | 10 | 10 | 9 | 9 | 46 | 11 | 11 | 10 | 8 | 8 | 6 | 54 |
| 2006 | 3 | 3 | 7 | 7 | 5 | 10 | 35 | 13 | 13 | 12 | 10 | 9 | 8 | 65 |
| 2007 | 1 | 4 | 5 | 6 | 6 | 8 | 30 | 9 | 10 | 10 | 9 | 9 | 10 | 57 |
| 2008 | - | - | 7 | 8 | 8 | 8 | 31 | 8 | 8 | 10 | 8 | 8 | 2 | 44 |
| 2009 | - | 5 | 6 | 8 | 8 | 9 | 36 | 7 | 9 | 10 | 9 | 7 | 2 | 44 |
| 2010 | 8 | 8 | 7 | 9 | 10 | 11 | 53 | 11 | 9 | 9 | 10 | 10 | 2 | 51 |
| 2011 | 4 | 8 | 8 | 10 | 8 | 7 | 45 | 7 | 8 | 8 | 9 | 4 | 1 | 37 |
| 2012 | 7 | 7 | 7 | 7 | 5 | 5 | 38 | 8 | 8 | 8 | 6 | 5 | 3 | 38 |
| 2013 | 4 | 6 | 6 | 7 | 5 | 6 | 34 | 6 | 6 | 6 | 7 | 7 | 7 | 39 |
| 2014 | 3 | 5 | 6 | 6 | 6 | 5 | 31 | 7 | 7 | 7 | 8 | 7 | 6 | 42 |
| 2015 | 1 | 4 | 6 | 6 | 6 | 7 | 30 | 7 | 7 | 8 | 8 | 8 | 8 | 46 |



Figure A.7. Standardised residuals for the Area 912 CRA 4 sub-stock standardised seasonal F2_LFX CPUE analysis.


Figure A.8. Coefficients for month from the Area 912 CRA 4 sub-stock seasonal F2_LFX CPUE standardisation. Month coefficients are not in canonical form, with each of the two reference months (September and October) set to $\mathbf{1 . 0}$ and the associated SE set to zero.


Standardised index error bars $=+/-1.96 *$ SE

Figure A.9. Scaled standardised F2_LFX CPUE (kg/potlift) by period for the Area 912 CRA 4 sub-stock with the AW-SS seasons plotted separately. Also shown are the arithmetic or "raw" CPUE series and the geometric mean of the CPUE ("unstandardised"). The standardised and unstandardised series were scaled by multiplying each index in the unscaled series (where the geometric mean=1) by the geometric mean of the arithmetic CPUE series for each seasonal category (geometric mean for $\mathrm{AW}=\mathbf{0 . 7 1} \mathrm{kg} / \mathrm{potlift}$; geometric mean for $\mathbf{S S}=\mathbf{1 . 0 0} \mathbf{~ k g} /$ potlift).

Table A.9. Total deviance ( $\mathrm{R}^{2}$ ) explained by each variable in the Area 912 CRA 4 sub-stock standardised seasonal CPUE model.

| Variable | $\mathbf{1}$ | $\mathbf{2}$ |
| :--- | ---: | ---: |
| Period | 0.3112 |  |
| Month | 0.0646 | 0.3596 |
| Additional deviance explained | 0.0000 | 0.0484 |

Table A.10: Standardised seasonal CPUE and standard errors for the Area 912 CRA 4 sub-stock.

| Fishing |  |  |  | Fishing |  |  |  |  |  |  |  | s.e. | SS | s.e. |
| :---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | :---: | :---: | :---: | :---: | :---: |
| Year | AW | s.e. | SS | S.e. | Year | AW | s.e. |  |  |  |  |  |  |  |
| 1979 | 0.796 | 0.0576 | 0.917 | 0.0478 | 1998 | 1.931 | 0.0612 | 3.298 | 0.1311 |  |  |  |  |  |
| 1980 | 1.054 | 0.0544 | 0.994 | 0.0459 | 1999 | 1.645 | 0.0632 | 2.291 | 0.1312 |  |  |  |  |  |
| 1981 | 1.110 | 0.0524 | 1.100 | 0.0464 | 2000 | 1.218 | 0.0611 | 2.794 | 0.1031 |  |  |  |  |  |
| 1982 | 0.905 | 0.0499 | 0.867 | 0.0462 | 2001 | 0.905 | 0.0623 | 1.496 | 0.1013 |  |  |  |  |  |
| 1983 | 0.686 | 0.0524 | 0.742 | 0.0476 | 2002 | 1.096 | 0.0618 | 1.208 | 0.0831 |  |  |  |  |  |
| 1984 | 0.722 | 0.0549 | 0.850 | 0.0509 | 2003 | 0.914 | 0.0653 | 1.230 | 0.0738 |  |  |  |  |  |
| 1985 | 0.648 | 0.0518 | 0.909 | 0.0482 | 2004 | 0.565 | 0.0658 | 0.902 | 0.0671 |  |  |  |  |  |
| 1986 | 0.633 | 0.0540 | 0.844 | 0.0492 | 2005 | 0.449 | 0.0786 | 0.705 | 0.0728 |  |  |  |  |  |
| 1987 | 0.548 | 0.0551 | 0.688 | 0.0503 | 2006 | 0.479 | 0.0893 | 0.632 | 0.0667 |  |  |  |  |  |
| 1988 | 0.426 | 0.0562 | 0.544 | 0.0536 | 2007 | 0.440 | 0.0963 | 0.723 | 0.0711 |  |  |  |  |  |
| 1989 | 0.425 | 0.0562 | 0.716 | 0.0515 | 2008 | 0.575 | 0.0950 | 0.858 | 0.0803 |  |  |  |  |  |
| 1990 | 0.567 | 0.0564 | 0.899 | 0.0518 | 2009 | 0.468 | 0.0883 | 0.850 | 0.0804 |  |  |  |  |  |
| 1991 | 0.612 | 0.0510 | 0.826 | 0.0494 | 2010 | 0.584 | 0.0734 | 0.922 | 0.0748 |  |  |  |  |  |
| 1992 | 0.505 | 0.0472 | 0.721 | 0.0465 | 2011 | 0.618 | 0.0795 | 1.057 | 0.0873 |  |  |  |  |  |
| 1993 | 0.392 | 0.0479 | 0.672 | 0.0547 | 2012 | 0.759 | 0.0863 | 0.945 | 0.0861 |  |  |  |  |  |
| 1994 | 0.617 | 0.0572 | 1.066 | 0.0785 | 2013 | 0.553 | 0.0909 | 0.835 | 0.0852 |  |  |  |  |  |
| 1995 | 0.846 | 0.0606 | 1.216 | 0.0829 | 2014 | 0.524 | 0.0951 | 0.699 | 0.0822 |  |  |  |  |  |
| 1996 | 1.397 | 0.0628 | 1.675 | 0.1122 | 2015 | 0.440 | 0.0964 | 0.850 | 0.0787 |  |  |  |  |  |
| 1997 | 1.885 | 0.0679 | 2.456 | 0.1659 |  |  |  |  |  |  |  |  |  |  |

## A.3.1.2 Area 913

The Area 913 data set shows a diminishing number of records over time (Table A.11). The total deviance explained by the Area 913 model was $39 \%$ (Table A.12), with only month available for standardisation apart from the time period variable. Residual patterns showed some deviation from the lognormal assumption at the peak of the residual distribution (Figure A.10).

The month categorical variable in the seasonal CPUE analysis appears to be cyclical, with a winter peak in May/June and a summer peak in December, but extending across November to January (Figure A.11). Both the Area 913 AW and SS CPUE series showed similar patterns, with the AW and SS series having approximately the same absolute catch rates, unlike the other three Stocks (Figure A.12, Table A.13). Both series peak twice: once in the late 1990s and a second peak around 2010 (Figure A.12). Both series have similar associated error, except during four years in the late 1990s, reflecting the small amount of data in the SS series in those years (Figure A.12).

Table A.11. Number of vessel/month records in the dataset used to calculate the Area 913 CRA 4 substock CPUE time series (based on the F2_LFX algorithm). '-': no data.



Figure A.10. Standardised residuals for the Area 913 CRA 4 sub-stock standardised seasonal F2_LFX CPUE analysis.


Figure A.11. Coefficients for month from the Area 913 CRA 4 sub-stock seasonal F2_LFX CPUE standardisation. Month coefficients are not in canonical form, with each of the two reference months (June and December) set to $\mathbf{1 . 0}$ and the associated SE set to zero.


Standardised index error bars=+/-1.96*SE

Figure A.12. Scaled standardised F2_LFX CPUE (kg/potlift) by period for the Area 913 CRA 4 sub-stock with the AW-SS seasons plotted separately. Also shown are the arithmetic or "raw" CPUE series and the geometric mean of the CPUE ("unstandardised"). The standardised and unstandardised series were scaled by multiplying each index in the unscaled series (where the geometric mean=1) by the geometric mean of the arithmetic CPUE series for each seasonal category (geometric mean for $\mathrm{AW}=\mathbf{0 . 9 2} \mathbf{~ k g} / \mathrm{potlift}$; geometric mean for $\mathbf{S S}=\mathbf{1 . 0 5} \mathbf{~ k g} /$ potlift).

Table A.12. Total deviance ( $\mathrm{R}^{2}$ ) explained by each variable in the Area 913 CRA 4 sub-stock standardised seasonal CPUE model.

| Variable | $\mathbf{1}$ | $\mathbf{2}$ |
| :--- | ---: | ---: |
| Period | 0.3233 |  |
| Month | 0.0737 | 0.3892 |
| Additional deviance explained | 0.0000 | 0.0659 |

Table A.13: Standardised seasonal CPUE and standard errors for the Area 913 CRA 4 sub-stock.

| Fishing |  |  | Fishing |  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | :---: | :---: |
| Year | AW | s.e. | SS | s.e. | Year | AW | s.e. | SS | s.e. |  |  |
| 1979 | 1.085 | 0.0611 | 1.006 | 0.0523 | 1998 | 2.751 | 0.0898 | 1.492 | 0.3004 |  |  |
| 1980 | 1.022 | 0.0614 | 0.683 | 0.0579 | 1999 | 1.920 | 0.0865 | 2.797 | 0.3671 |  |  |
| 1981 | 1.057 | 0.0642 | 0.753 | 0.0628 | 2000 | 2.133 | 0.0952 | 2.373 | 0.1656 |  |  |
| 1982 | 1.100 | 0.0668 | 0.900 | 0.0559 | 2001 | 1.330 | 0.0824 | 1.713 | 0.1176 |  |  |
| 1983 | 1.111 | 0.0599 | 0.871 | 0.0529 | 2002 | 1.283 | 0.0659 | 1.766 | 0.1018 |  |  |
| 1984 | 0.826 | 0.0578 | 0.802 | 0.0563 | 2003 | 1.397 | 0.0631 | 1.755 | 0.0833 |  |  |
| 1985 | 0.687 | 0.0622 | 0.836 | 0.0617 | 2004 | 1.106 | 0.0720 | 1.434 | 0.0774 |  |  |
| 1986 | 0.767 | 0.0595 | 0.902 | 0.0545 | 2005 | 0.979 | 0.0738 | 0.851 | 0.0781 |  |  |
| 1987 | 0.667 | 0.0611 | 0.794 | 0.0569 | 2006 | 0.682 | 0.0737 | 0.832 | 0.0661 |  |  |
| 1988 | 0.595 | 0.0644 | 0.684 | 0.0573 | 2007 | 0.635 | 0.0983 | 0.734 | 0.0767 |  |  |
| 1989 | 0.581 | 0.0579 | 0.828 | 0.0529 | 2008 | 0.613 | 0.0997 | 0.865 | 0.0786 |  |  |
| 1990 | 0.451 | 0.0594 | 0.570 | 0.0558 | 2009 | 0.951 | 0.0843 | 1.154 | 0.0896 |  |  |
| 1991 | 0.518 | 0.0546 | 0.591 | 0.0542 | 2010 | 0.969 | 0.0841 | 1.315 | 0.0876 |  |  |
| 1992 | 0.374 | 0.0538 | 0.510 | 0.0525 | 2011 | 1.184 | 0.0720 | 1.420 | 0.0897 |  |  |
| 1993 | 0.515 | 0.0549 | 0.628 | 0.0556 | 2012 | 1.528 | 0.0833 | 1.836 | 0.0852 |  |  |
| 1994 | 0.573 | 0.0539 | 0.717 | 0.0608 | 2013 | 1.402 | 0.0951 | 1.303 | 0.0929 |  |  |
| 1995 | 0.680 | 0.0626 | 1.147 | 0.0768 | 2014 | 0.966 | 0.0936 | 0.919 | 0.0832 |  |  |
| 1996 | 1.010 | 0.0740 | 1.477 | 0.1749 | 2015 | 0.731 | 0.0936 | 0.950 | 0.0758 |  |  |
| 1997 | 1.519 | 0.0833 | 2.403 | 0.2601 |  |  |  |  |  |  |  |

## A.3.1.3 Area 914

As with the other CRA 4 sub-stocks, the Area 914 data set shows a diminishing number of records over time (Table A.14). The total deviance explained by the Area 914 model was $43 \%$ (Table A.15), with only month available for standardisation apart from the time period variable. Residual patterns showed some deviation from the lognormal assumption at the peak of the residual distribution (Figure A.13).

As for the other CRA 4 sub-stocks, the month categorical variable in the seasonal CPUE analysis appears to be cyclical, with a winter peak in May/June and a summer peak in November/December (Figure A.14). Both the Area 914 AW and SS CPUE series showed similar patterns, but the first peak is lower than the second peak for the AW series, and the AW series has lower absolute catch rates than the SS series (Figure A.15, Table A.16). Both series peak twice: once in the late 1990s and a second peak around 2012-13 (Figure A.15). The associated error is greater for the SS series in both of the peak years, reflecting small amounts of data in the SS series in those years (Figure A.15).

Table A.14. Number of vessel/month records in the dataset used to calculate the Area 914 CRA 4 substock CPUE time series (based on the F2_LFX algorithm). ' - ': no data.



Figure A.13. Standardised residuals for the Area 914 CRA 4 sub-stock standardised seasonal F2_LFX CPUE analysis.


Figure A.14. Coefficients for month from the Area 914 CRA 4 sub-stock seasonal F2_LFX CPUE standardisation. Month coefficients are not in canonical form, with each of the two reference months (July and October) set to $\mathbf{1 . 0}$ and the associated SE set to zero.


Figure A.15. Scaled standardised F2_LFX CPUE (kg/potlift) by period for the Area 914 CRA 4 sub-stock with the AW-SS seasons plotted separately. Also shown are the arithmetic or "raw" CPUE series and the geometric mean of the CPUE ("unstandardised"). The standardised and unstandardised series were scaled by multiplying each index in the unscaled series (where the geometric mean=1) by the geometric mean of the arithmetic CPUE series for each seasonal category (geometric mean for AW=0.74 kg/potlift; geometric mean for $\mathbf{S S}=\mathbf{1 . 0 4} \mathbf{~ k g} /$ potlift).

Table A.15. Total deviance ( $\mathrm{R}^{2}$ ) explained by each variable in the Area 914 CRA 4 sub-stock standardised seasonal CPUE model.

| Variable | $\mathbf{1}$ | $\mathbf{2}$ |
| :--- | ---: | ---: |
| Period | 0.3460 |  |
| Month | 0.0920 | 0.4313 |
| Additional deviance explained | 0.0000 | 0.0853 |

Table A.16: Standardised seasonal CPUE and standard errors for the Area 914 CRA 4 sub-stock.

| Fishing |  |  | Fishing |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | :---: |
| Year | AW | s.e. | SS | s.e. | Year | AW | s.e. | SS | s.e. |  |
| 1979 | 0.785 | 0.0534 | 0.793 | 0.0449 | 1998 | 1.049 | 0.0493 | 2.432 | 0.1288 |  |
| 1980 | 0.810 | 0.0510 | 0.715 | 0.0451 | 1999 | 1.039 | 0.0506 | 1.578 | 0.1117 |  |
| 1981 | 0.649 | 0.0507 | 0.818 | 0.0473 | 2000 | 0.976 | 0.0519 | 1.913 | 0.1176 |  |
| 1982 | 1.105 | 0.0478 | 1.184 | 0.0422 | 2001 | 0.967 | 0.0513 | 1.829 | 0.0917 |  |
| 1983 | 1.168 | 0.0471 | 0.968 | 0.0422 | 2002 | 0.866 | 0.0488 | 2.117 | 0.0739 |  |
| 1984 | 0.843 | 0.0446 | 0.737 | 0.0430 | 2003 | 0.932 | 0.0501 | 1.903 | 0.0729 |  |
| 1985 | 0.597 | 0.0449 | 0.801 | 0.0420 | 2004 | 0.716 | 0.0487 | 1.456 | 0.0591 |  |
| 1986 | 0.744 | 0.0454 | 0.942 | 0.0428 | 2005 | 0.867 | 0.0561 | 0.989 | 0.0538 |  |
| 1987 | 0.582 | 0.0464 | 0.859 | 0.0418 | 2006 | 0.477 | 0.0555 | 0.705 | 0.0481 |  |
| 1988 | 0.535 | 0.0460 | 0.689 | 0.0431 | 2007 | 0.366 | 0.0616 | 0.566 | 0.0541 |  |
| 1989 | 0.456 | 0.0463 | 0.566 | 0.0432 | 2008 | 0.470 | 0.0729 | 0.745 | 0.0689 |  |
| 1990 | 0.337 | 0.0499 | 0.481 | 0.0445 | 2009 | 0.970 | 0.0736 | 1.384 | 0.0812 |  |
| 1991 | 0.312 | 0.0443 | 0.492 | 0.0449 | 2010 | 0.903 | 0.0597 | 1.401 | 0.0622 |  |
| 1992 | 0.385 | 0.0465 | 0.508 | 0.0474 | 2011 | 1.226 | 0.0556 | 1.567 | 0.0653 |  |
| 1993 | 0.510 | 0.0443 | 0.765 | 0.0555 | 2012 | 1.349 | 0.0562 | 1.813 | 0.0625 |  |
| 1994 | 0.661 | 0.0409 | 1.019 | 0.0664 | 2013 | 1.201 | 0.0548 | 1.598 | 0.0581 |  |
| 1995 | 0.822 | 0.0414 | 1.261 | 0.0775 | 2014 | 1.011 | 0.0547 | 1.202 | 0.0535 |  |
| 1996 | 1.034 | 0.0470 | 1.604 | 0.1248 | 2015 | 0.596 | 0.0548 | 0.735 | 0.0491 |  |
| 1997 | 1.020 | 0.0452 | 1.393 | 0.1383 |  |  |  |  |  |  |

## A.3.1.4 Area 915+934

Again, as with the other CRA 4 sub-stocks, the Area 915+934 data set shows a diminishing number of records over time (Table A.17). The total deviance explained by the Area $915+934$ model was $37 \%$ (Table A.18), with area having slightly more explanatory power over month in the standardisation procedure. Residual patterns showed some deviation from the lognormal assumption at the peak of the residual distribution (Figure A.16).

The month categorical variable in the seasonal CPUE has a winter peak in June and there is no summer peak (Figure A.17, right panel). The coefficient for Area 934 exceeds that of the Area 915 coefficient (Figure A.17, left panel). Both the Area 915+934 AW and SS CPUE series showed similar patterns, with a suggestion that there might be an initial peak in the late 1990s in the SS series, although the model uncertainty is high (Figure A.18, Table A.19). Absolute catch rates are higher in the SS series and there is no AW peak in the 1990s (Figure A.18). The associated error is greater for the SS series, reflecting smaller amounts of data in the SS series (Figure A.18).

Table A.17. Number of vessel/month records in the dataset used to calculate the Area 915+934 CRA 4 sub-stock CPUE time series (based on the F2_LFX algorithm). '-': no data.



Figure A.16. Standardised residuals for the Area 915+934 CRA 4 sub-stock standardised seasonal F2_LFX CPUE analysis.


Figure A.17. Coefficients for statistical area and month from the Area 915+934 CRA 4 sub-stock seasonal F2_LFX CPUE standardisation. Month coefficients are not in canonical form, with each of the two reference months (September and October) set to 1.0 and the associated SE set to zero.


Standardised index error bars $=+/-1.96 *$ SE

Figure A.18. Scaled standardised F2_LFX CPUE (kg/potlift) by period for the Area 915+934 CRA 4 substock with the AW-SS seasons plotted separately. Also shown are the arithmetic or "raw" CPUE series and the geometric mean of the CPUE ("unstandardised"). The standardised and unstandardised series were scaled by multiplying each index in the unscaled series (where the geometric mean=1) by the geometric mean of the arithmetic CPUE series for each seasonal category (geometric mean for AW=0.62 kg/potlift; geometric mean for $\mathrm{SS}=\mathbf{0 . 8 7} \mathbf{~ k g} /$ potlift).

Table A.18. Total deviance ( $\mathrm{R}^{\mathbf{2}}$ ) explained by each variable in the Area $915+934$ CRA 4 sub-stock standardised seasonal CPUE model.

| Variable | $\mathbf{1}$ | $\mathbf{2}$ | $\mathbf{3}$ |
| :--- | ---: | ---: | ---: |
| Period | 0.3233 |  |  |
| Area | 0.0456 | 0.3459 |  |
| Month | 0.0333 | 0.3458 | 0.3682 |
| Additional deviance explained | 0.0000 | 0.0226 | 0.0222 |

Table A.19: Standardised seasonal CPUE and standard errors for the Area 915+934 CRA 4 sub-stock.

| Fishing |  |  |  | Fishing |  |  |  |  |  |
| :---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Year | AW | s.e. | SS | s.e. | Year | AW | s.e. | SS | s.e. |
| 1979 | 0.535 | 0.0772 | 0.649 | 0.0582 | 1998 | 0.729 | 0.0867 | 1.097 | 0.1562 |
| 1980 | 0.450 | 0.0712 | 0.641 | 0.0538 | 1999 | 0.774 | 0.0872 | 1.340 | 0.1906 |
| 1981 | 0.651 | 0.0752 | 0.627 | 0.0585 | 2000 | 0.692 | 0.0826 | 1.074 | 0.1350 |
| 1982 | 0.545 | 0.0695 | 0.667 | 0.0554 | 2001 | 0.730 | 0.0789 | 0.970 | 0.1800 |
| 1983 | 0.557 | 0.0648 | 0.601 | 0.0564 | 2002 | 0.585 | 0.0885 | 1.674 | 0.1137 |
| 1984 | 0.558 | 0.0648 | 0.559 | 0.0587 | 2003 | 0.718 | 0.0881 | 1.640 | 0.0940 |
| 1985 | 0.547 | 0.0663 | 0.629 | 0.0566 | 2004 | 0.601 | 0.0946 | 1.345 | 0.1054 |
| 1986 | 0.503 | 0.0685 | 0.649 | 0.0601 | 2005 | 0.651 | 0.0989 | 0.855 | 0.0914 |
| 1987 | 0.443 | 0.0766 | 0.640 | 0.0632 | 2006 | 0.722 | 0.0785 | 0.848 | 0.0665 |
| 1988 | 0.375 | 0.0745 | 0.567 | 0.0693 | 2007 | 0.642 | 0.0864 | 0.703 | 0.0737 |
| 1989 | 0.345 | 0.0832 | 0.490 | 0.0763 | 2008 | 0.846 | 0.0998 | 1.136 | 0.0988 |
| 1990 | 0.391 | 0.0791 | 0.470 | 0.0679 | 2009 | 1.234 | 0.1002 | 1.542 | 0.0878 |
| 1991 | 0.347 | 0.0841 | 0.453 | 0.0741 | 2010 | 0.873 | 0.0765 | 1.265 | 0.0808 |
| 1992 | 0.406 | 0.0824 | 0.527 | 0.0747 | 2011 | 1.212 | 0.0956 | 1.749 | 0.0935 |
| 1993 | 0.348 | 0.0746 | 0.491 | 0.0800 | 2012 | 1.144 | 0.1019 | 1.589 | 0.0906 |
| 1994 | 0.402 | 0.0724 | 0.536 | 0.0902 | 2013 | 0.942 | 0.1137 | 1.539 | 0.1115 |
| 1995 | 0.561 | 0.0866 | 0.726 | 0.1352 | 2014 | 1.277 | 0.0970 | 1.823 | 0.0922 |
| 1996 | 0.675 | 0.0789 | 1.161 | 0.2196 | 2015 | 0.783 | 0.1000 | 1.162 | 0.0838 |
| 1997 | 0.838 | 0.0942 | 0.802 | 0.2198 |  |  |  |  |  |

## A.3.2 Historical Catch Rate (CR) Data

Catch and effort (days fishing) data from 1963 through 1973 from the Annala \& King (1983) data set were allocated to rock lobster statistical areas using the procedure described in Section A.2.1. These data were used to calculate unstandardised catch per day for each calendar year from 1963 to 1973 for the four defined stocks (Table A.20) which are plotted after being normalised to the same geometric mean (Figure A.19). Note that the series for Areas 913 and 914 are the same because they have been derived from the same Annala \& King statistical area (see Figure A.2).

Table A.20: Catch rate (kg/day) from the potlift data in Annala \& King (1983) calculated for each statistical area after applying the algorithm described in Steps 1 to 5 (above). The calculated $p k_{y}^{a}$ was also applied to the days fishing field before calculating the annual catch rate. The total CRA 4 catch rate values are those that are stored in CRACE.

| Calendar | Stock/statistical area |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: |
| year | $\mathbf{9 1 2}$ | $\mathbf{9 1 3}$ | $\mathbf{9 1 4}$ | $\mathbf{9 1 5 + 9 3 4}$ | CRA 4 |
| 1963 | 145.7 | 304.4 | 304.4 | 121.1 | 69.2 |
| 1964 | 211.8 | 409.9 | 409.9 | 152.3 | 102.2 |
| 1965 | 285.4 | 377.2 | 377.2 | 129.3 | 114.5 |
| 1966 | 245.3 | 328.8 | 328.8 | 148.7 | 108.0 |
| 1967 | 238.2 | 245.8 | 245.8 | 143.9 | 93.6 |
| 1968 | 162.4 | 193.2 | 193.2 | 103.1 | 69.1 |
| 1969 | 157.6 | 137.3 | 137.3 | 99.7 | 60.1 |
| 1970 | 115.9 | 141.3 | 141.3 | 88.6 | 53.2 |
| 1971 | 95.6 | 136.3 | 136.3 | 96.7 | 50.4 |
| 1972 | 77.9 | 143.6 | 143.6 | 89.1 | 46.8 |
| 1973 | 103.9 | 136.0 | 136.0 | 70.1 | 47.8 |

Annala\&King catch rate: CRA4 \& sub-stock


Figure A.19: Annala \& King catch rates (Table A.20), normalised to the geometric mean, plotted for the four CRA 4 sub-stocks and the overall CRA 4 series.

## A. 4 Length frequency data

The distribution of length frequencies for each of the five statistical areas that comprise the CRA 4 sub-stock are plotted by fishing year and season for males (Figure A.20) and females (Figure A.21). The distributions of the immature females are superimposed in black in the female plots. These plots show the availability of data in each of the four CRA 4 multi-area sub-stocks. While there are only very small amounts of data available for Area 934, they are largely consistent with the Area 915 data, justifying the amalgamation of these two statistical areas into a single region.
Table A. 21 shows the number of lobsters measured by CRA 4 multi-area substock, year, season and catch sampling source.

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Figure A.20: Male length frequency distributions ( $\mathbf{3 0}-\mathbf{9 0} \mathbf{m m}$ ) by Statistical Area, year and season. The number of males sampled is given in the upper right side of the figures. The current MLS of 54 mm is shown with a vertical blue line.




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Figure A. 20 (cont.): Male length frequency distributions ( $\mathbf{3 0 - 9 0} \mathbf{~ m m}$ ) by Statistical Area, year and season. The number of males sampled is given in the upper right side of the figures. The current MLS of 54 mm is shown with a vertical blue line.

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Figure A. 20 (cont.): Male length frequency distributions ( $\mathbf{3 0 - 9 0} \mathbf{~ m m}$ ) by Statistical Area, year and season. The number of males sampled is given in the upper right side of the figures. The current MLS of 54 mm is shown with a vertical blue line.


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Figure A. 20 (cont.): Male length frequency distributions ( $\mathbf{3 0 - 9 0} \mathbf{~ m m}$ ) by Statistical Area, year and season. The number of males sampled is given in the upper right side of the figures. The current MLS of 54 mm is shown with a vertical blue line.


Figure A. 20 (cont.): Male length frequency distributions ( $30-90 \mathrm{~mm}$ ) by Statistical Area, year and season. The number of males sampled is given in the upper right side of the figures. The current MLS of 54 mm is shown with a vertical blue line


Figure A. 20 (cont.): Male length frequency distributions ( $\mathbf{3 0 - 9 0} \mathbf{~ m m}$ ) by Statistical Area, year and season. The number of males sampled is given in the upper right side of the figures. The current MLS of 54 mm is shown with a vertical blue line


Figure A. 20 (cont.): Male length frequency distributions ( $\mathbf{3 0 - 9 0} \mathbf{~ m m}$ ) by Statistical Area, year and season. The number of males sampled is given in the upper right side of the figures. The current MLS of 54 mm is shown with a vertical blue line.


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Figure A.21: Female length frequency distributions ( $\mathbf{3 0 - 9 0} \mathbf{~ m m}$ ) by Statistical Area, year and season. Immature females are shown in black and mature females in grey. The number of females sampled is given in the upper right side of the figures. The current MLS of $\mathbf{6 0} \mathbf{m m}$ is shown with a vertical blue line.

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Figure A. 21 (cont.): Female length frequency distributions ( $\mathbf{3 0}-\mathbf{9 0} \mathbf{~ m m}$ ) by Statistical Area, year and season. Immature females are shown in black and mature females in grey. The number of females sampled is given in the upper right side of the figures. The current MLS of $\mathbf{6 0} \mathbf{~ m m}$ is shown with a vertical blue line.

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Figure A. 21 (cont.): Female length frequency distributions ( $30-90 \mathrm{~mm}$ ) by Statistical Area, year and season. Immature females are shown in black and mature females in grey. The number of females sampled is given in the upper right side of the figures. The current MLS of $\mathbf{6 0} \mathbf{~ m m}$ is shown with a vertical blue line.


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Figure A. 21 (cont.): Female length frequency distributions ( $\mathbf{3 0} \mathbf{- 9 0} \mathbf{~ m m}$ ) by Statistical Area, year and season. Immature females are shown in black and mature females in grey. The number of females sampled is given in the upper right side of the figures. The current MLS of $\mathbf{6 0} \mathbf{~ m m}$ is shown with a vertical blue line.


Figure A. 21 (cont.): Female length frequency distributions ( $\mathbf{3 0} \mathbf{- 9 0} \mathbf{~ m m}$ ) by Statistical Area, year and season. Immature females are shown in black and mature females in grey. The number of females sampled is given in the upper right side of the figures. The current MLS of $\mathbf{6 0} \mathbf{~ m m}$ is shown with a vertical blue line.


Figure A. 21 (cont.): Female length frequency distributions ( $\mathbf{3 0} \mathbf{- 9 0} \mathbf{~ m m}$ ) by Statistical Area, year and season. Immature females are shown in black and mature females in grey. The number of females sampled is given in the upper right side of the figures. The current MLS of $\mathbf{6 0} \mathbf{~ m m}$ is shown with a vertical blue line.


Figure A. 21 (cont.): Female length frequency distributions ( $\mathbf{3 0} \mathbf{- 9 0} \mathbf{m m}$ ) by Statistical Area, year and season. Immature females are shown in black and mature females in grey. The number of females sampled is given in the upper right side of the figures. The current MLS of 60 mm is shown with a vertical blue line.

Table A.21: Number of lobsters measured by CRA 4 sub-stock, season and catch sampling source. codes: LB=logbook; CS=catch sampling.

|  | 912 |  |  |  | 913 |  |  |  | 914 |  |  |  | 915+934 |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Fishing |  | AW |  | SS |  | AW |  | SS |  | AW |  | SS |  | AW |  | SS |
| year | LB | CS | LB | CS | LB | CS | LB | CS | LB | CS | LB | CS | LB | CS | LB | CS |
| 1986 | - | - | - | 276 | - | - | - | - | - | - | - | - | - | - | - | - |
| 1987 | - | 1194 | - | 1564 | - | - | - | - | - | - | - | - | - | - | - | - |
| 1988 | - | 1980 | - | 1851 | - | - | - | - | - | - | - | - | - | - | - | - |
| 1989 | - | 897 | - | 2330 | - | 1092 | - | 1492 | - | - | - | - | - | 1672 | - | 791 |
| 1990 | - | 3171 | - | 10117 | - | 874 | - | 5868 | - | - | - | - | - | 3806 | - | 1185 |
| 1991 | - | 2984 | - | 10904 | - | - | - | 4751 | - | - | - | - | - | - | - | - |
| 1992 | - | 1502 | - | 13914 | - | - | - | 2632 | - | - | - | - | - | - | - | - |
| 1993 | - | 1112 | - | 7775 | - | - | - | 3016 | - | - | - | - | - | - | - | - |
| 1994 | - | 2540 | - | 4415 | - | - | - | 1115 | - | - | - | - | - | - | - | - |
| 1995 | - | 2395 | - | 5909 | - | - | - | 1464 | - | - | - | - | - | - | - | - |
| 1996 | - | 2434 | - | - | - | 1010 | - | - | - | 1105 | - | - | - | - | - | - |
| 1997 | - | 14252 | - | - | - | 10217 | - | - | 1774 | 9532 | 70 | - | - | - | - | - |
| 1998 | - | 6275 | - | - | - | 9388 | - | - | 811 | 3469 | 586 | - | - | - | - | - |
| 1999 | - | 11294 | - | - | - | 7516 | - | - | 297 | 4868 | - | - | - | 1437 | - | - |
| 2000 | - | 6379 | - | - | - | 7461 | - | - | - | 5553 | - | - | 331 | 3131 | - | - |
| 2001 | - | 6934 | - | - | - | 4724 | - | 1983 | - | 4494 | - | 1394 | - | 1005 | - | 184 |
| 2002 | - | 3796 | - | 1172 | - | 6604 | - | 538 | - | 7105 | - | 1171 | 494 | 591 | 98 | - |
| 2003 | - | 3733 | - | - | - | 5030 | - | 1350 | - | 5183 | - | 1422 | 1252 | 1176 | 246 | 664 |
| 2004 | - | 2556 | - | 463 | - | 5595 | - | 5278 | - | 4483 | - | 4108 | 935 | 603 | 88 | - |
| 2005 | - | 1762 | - | 1289 | 395 | 3262 | - | 7014 | - | 4421 | - | 6445 | 1766 | - | 300 | 1890 |
| 2006 | - | 286 | - | 615 | 196 | 1389 | - | 3986 | - | 5292 | - | 6610 | 1103 | 1668 | 107 | 2492 |
| 2007 | - | 249 | - | 988 | - | 1216 | - | 4799 | - | 4386 | - | 4416 | 738 | 1568 | 662 | 2684 |
| 2008 | - | - | - | 2187 | - | - | - | 3564 | - | 2356 | - | 4772 | 522 | 1967 | - | 5460 |
| 2009 | - | 557 | - | 4589 | - | 1786 | - | 2303 | - | 2304 | 37 | 6482 | 865 | 1513 | 43 | 796 |
| 2010 | - | 721 | - | 1564 | 645 | 2011 | 804 | 828 | 506 | 3220 | 264 | 6537 | 1055 | 2157 | 396 | 1302 |
| 2011 | - | 1582 | - | 820 | 1742 | 2879 | 699 | 1937 | 1967 | 5933 | 1233 | 4099 | 31 | 1238 | - | 802 |
| 2012 | - | 1128 | - | 449 | 1186 | 4658 | 1356 | 4550 | 1393 | 5860 | 496 | 3380 | 226 | 1162 | 1329 | 648 |
| 2013 | - | 259 | - | 1478 | 893 | 2407 | 406 | 5973 | 201 | 5401 | 235 | - | 596 | 308 | 1227 | 1049 |
| 2014 | - | - | 216 | 699 | 703 | 3544 | 1605 | 1813 | 1060 | 4169 | 1664 | 3988 | 1411 | 1190 | 1946 | 925 |
| 2015 | 29 | - | 400 | 788 | 1382 | 592 | 2351 | 2259 | 1143 | 4767 | 1702 | 5302 | 1506 | 606 | 1656 | 704 |

## A. 5 TAGging data

The following three tables (Table A.22, Table A.23, Table A.24) provide basic information on the availability of tagging data for each of the four CRA 4 sub-stocks.

Table A.22: Number of tag recoveries by sex and Statistical Area of release for CRA 4 tag releases.

| Area | Male | Female | Total |
| :--- | ---: | ---: | ---: |
| 912 | 333 | 52 | 385 |
| 913 | 655 | 78 | 733 |
| 914 | 695 | 285 | 980 |
| 915 | 155 | 215 | 370 |
| Total | 1838 | 630 | 2468 |

Table A.23: Number of tag recoveries by year and area of release.

| Year | $\mathbf{9 1 2 + 9 1 3}$ | $\mathbf{9 1 4}+\mathbf{9 1 5}$ | Total |
| :--- | ---: | ---: | ---: |
| 1982 | 10 |  | 10 |
| 1998 | 325 | 381 | 706 |
| 1999 | 255 | 207 | 462 |
| 2000 | 96 | 80 | 176 |
| 2001 | 4 | 1 | 5 |
| 2002 |  | 4 | 4 |
| 2003 |  | 6 | 6 |
| 2004 |  | 2 | 2 |
| 2005 | 21 | 169 | 190 |
| 2006 | 16 | 78 | 94 |
| 2007 | 16 | 132 | 148 |
| 2008 | 3 | 1 | 4 |
| 2009 | 32 | 42 | 74 |
| 2010 | 275 | 8 | 283 |
| 2011 | 50 | 2 | 52 |
| 2012 | 6 |  | 6 |
| 2014 | 9 | 234 | 243 |
| 2015 |  | 3 | 3 |
| Total | 1118 | 1350 | 2468 |

Table A.24: Number of tag recoveries by Statistical Area of release and of recovery.

| Area | of | Statistical Area of Recovery |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| release |  | $\mathbf{9 1 2}$ | $\mathbf{9 1 3}$ | $\mathbf{9 1 4}$ | $\mathbf{9 1 5}$ | $\mathbf{9 1 6}$ | $\mathbf{9 3 4}$ | Total


[^0]:    ${ }^{1}$ The fishing year runs from 1 April through 31 March; the fishing year is named by the April-December portion; viz. 2015-16 is called "2015".

